Visualizing Earnings to Predict Post-Earnings Announcement Drift: A Deep Learning Approach

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Abstract

We study the drift-predicting information contained in visualized earnings, by plotting firms' time series of quarterly earnings in bar charts and employing a deep learning algorithm. We use the convolutional neural network (CNN) to extract features that are most predictive of post-earnings announcement drift from these images. In out-of-sample tests, these features significantly predict post-earnings announcement returns. The predictability is incremental to that of the usual drift determinants, exhibits time stability, cannot be explained by a battery of risk controls or return anomalies, appears to be consistent with investors missing predictable implications of these features for future earnings growth, and is robust to alternative modeling choices.

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1 Introduction

Human brains process visual information faster than textual or numerical information, with visual stimuli being recognized and interpreted almost instantaneously. In addition, visualizations such as graphs, charts, and infographics can simplify complex data and reveal unanticipated patterns that are not readily apparent in a table of numbers. For example, earnings charts are usually presented during firms' earnings call conferences to showcase earnings performance in the past few quarters, and are widely used by investors and analysts to evaluate firms' future prospects. In this study, we examine whether one can extract relevant information from visualized earnings data to predict post-earnings announcement returns, with the help of artificial intelligence.

We transform firms' historical quarterly earnings into bar charts and apply convolutional neural network (CNN), a deep learning algorithm inspired by the human biological visual system, to extract features that are most predictive of firms' post-announcement performance. CNN is capable of extracting features in a hierarchical manner, with early layers capturing local features and deeper layers integrating these features to detect more complex and global patterns. Our use of CNN closely resembles the image classification task in other CNN applications. For example, CNN can be trained with many bird, cat, and dog images to *automatically* learn features that best distinguish between the three animals.¹ Then, one

¹In particular, CNN learns to map the input images to their correct labels (bird, cat, or dog) by adjusting its internal parameters (weights) to minimize the difference between its predictions and the actual labels. We elaborate more on how CNN works, in Section 3.

can apply the CNN stored parameters from the previous training phase, to a new image to generate each label's probability, which represents CNN-predicted likelihood for this image to be a bird, cat, or dog image, respectively.

In our context, we are interested in examining whether CNN can be employed to learn features containing drift-predicting information from visualized earnings data. To begin with, for each firm announcing quarterly earnings from 1974Q1 to 2023Q2, we plot its most recent eight quarterly earnings in a bar chart that visualizes the magnitude as well as the sign of the earnings. Each earnings bar-chart image is paired with one of the three labels ("sell", "hold", or "buy") based on the relative performance of the firm's 63-day post-announcement buy-and-hold abnormal returns among the cross-section of firms in the same quarter. We then train the CNN model with 124,413 earnings images in a 20-year in-sample period (1974Q1 to 1993Q4) to *automatically* learn features that best distinguish between the three assigned labels.

Next, we apply the CNN stored parameters from the previous training phase to each earnings image after 1993 to generate the "*CNN buy* probability". This is our key independent variable, and can be thought of as the CNN-predicted likelihood for an image to be a "buy" image *when "sell" and "hold" options are available*. We seek to use this variable to predict the next 63 days of BHARs (i.e. post-earnings announcement drift). If certain features in earnings images are strongly associated with post-earnings announcement drift, and CNN is capable of detecting these features (which we label *CNN buy* features) during the training process, then firms with higher *CNN buy* probability should experience higher post-earnings announcement drift. To test this, in each quarter we sort firms into decile portfolios based on their $CNN \ buy$ probability², and then examine the average 63-day post-announcement buy-and-hold abnormal returns for each decile portfolio in the out-of-sample period from 1994Q3 to 2023Q2.

We find that when moving from the lowest decile to the highest decile of CNN buy probability, the average 63-day post-announcement buy-and-hold abnormal returns monotonically increase. Importantly, firms in the highest CNN buy probability decile outperform firms in the lowest CNN buy probability decile by 3.6% (t-statistic = 7.213) using the market-adjusted buy-and-hold returns, with the return differential being positive in 99 out of 116 quarters. Also importantly, the lowest CNN buy decile associates with significantly negative BHARs, creating the more typical view of post-earnings announcement drift (positive surprises followed by positive drift and negative surprises followed by negative drift). This offers a clearer link to the original drift-puzzle than many of the follow-on papers which had difficulty documenting the negative drift part of the picture (e.g. Garfinkel and Sokobin, 2006).

Our results are robust to alternative measures of buy-and-hold abnormal returns, such as size-adjusted buy-and-hold returns or buy-and-hold returns adjusted by factor models including the Fama-French four- and six-factor models (Fama and French, 1993; Carhart, 1997; Fama and French, 2015; Fama and French, 2018), the q^5 -model (Hou et al., 2015; Hou et al., 2021), and the risk-and-behavioral model (Daniel et al., 2020). Hence, the CNN model's ability to detect drift-predicting features from earnings images cannot be attributed to existing factors.

We next compare the drift-predicting power of CNN buy features to those of the well-known

 $^{^2 {\}rm where \ the \ cutoffs}$ are based on the distribution of the previous quarter's $C\!N\!N \ buy$ probability to prevent hindsight bias

determinants of post-earnings announcement drift, including the standardized unexpected earnings (Ball and Brown, 1968; Bernard and Thomas, 1989; Foster et al., 1984), earnings acceleration (He and Narayanamoorthy, 2020), trend in gross profitability (Akbas et al., 2017), market capitalization (Fama and French, 1992, 1993), book-to-market ratio (Fama and French, 1992, 1993), earnings announcement return (Foster et al., 1984; Chan et al., 1996), pre-announcement return (Carhart, 1997), earnings persistence (Francis et al., 2004), earnings volatility (Cao and Narayanamoorthy, 2012), gross profitability (Novy-Marx, 2013), operating profitability (Ball et al., 2016), operating accruals (Sloan, 1996; Hribar and Collins, 2002), total accruals (Richardson et al., 2005), and asset growth (Cooper et al., 2008). Using univariate portfolio analysis, we find the post-announcement return differential between the highest and lowest deciles sorted on each of the 14 firm characteristics is smaller in magnitude compared to that sorted on the *CNN buy* probability.

We proceed to examine whether the drift-predicting power of the CNN buy features shares resemblance with that of the firm characteristics mentioned above. In two-way 5×5 sorting analysis, we find that the drift-predicting ability of the CNN buy features persists in the middle to highest SUE quintiles as well as all quintiles of the other firm characteristics. This suggests that CNN buy features are mostly distinct from existing firm characteristics known to predict drift. Then we simultaneously control for these other explanators of drift, by running quarterly weighted Fama and MacBeth (1973) regressions of post-earnings announcement drift on the CNN buy probability and the 14 firm characteristics. The coefficient on the CNN buy probability is positive and highly significant, indicating that the CNN buy features provides incremental predictability for post-earnings announcement drift when controlling for well-known extant determinants. Since the positive relation between *CNN buy* probability and post-earnings announcement drift cannot be explained by risk and known anomalies, we proceed to explore the nature of *CNN buy* features and the source of their drift-predicting power. In particular, we correlate *CNN buy* probability with several firm characteristics that have each shown prior evidence of return predictability. This suggests that the CNN model is capable of discerning meaningful return-predicting information solely from historical earnings represented in the form of an image.

In addition, we show that the non-linear transformations of CNN impart explanatory power. In particular, a regression of the CNN buy probability on the most recent eight quarterly earnings shows that historical earnings collectively explain 30% of the variation in CNN buy probability, suggesting that most of the variation in CNN buy probability is driven by nonlinear transformations of the underlying historical earnings data.

We then explore the nature of the price relevant information in *CNN buy* features that is apparently missed by the market, leading to the post-earnings announcement drift. We hypothesize that the *CNN buy* features have implications for future earnings growth and investors do not fully incorporate these implications in a timely fashion. This interpretation is in line with Bernard and Thomas, 1989, but is also complementary given the importance of non-linear transformations that CNN allows. Moreover, to the extent that arbitrageurs were expected to eliminate the original trading rules built on Bernard and Thomas, 1989, our results offer potential explanation for the persistence of the drift phenomenon.

As preliminary evidence, we find that *CNN buy* probability positively predicts one-quarter-ahead earnings growth as well as three-day abnormal returns around the next earnings announcement

date, controlling for past earnings growth and other firm characteristics.³ Then we conduct a formal market efficiency test, the Mishkin test (Mishkin, 1983; Abel and Mishkin, 1983), to show that the drift-predicting ability of CNN buy features likely manifests because investors underestimate the implications of CNN buy features for future earnings growth.⁴

We perform a battery of additional analyses as robustness checks. First, we find that employing CNN predictions in a more conservative monthly-rebalancing long-short strategy yields a monthly return of around 1%. Second, we show that the out-of-sample performance of CNN predictions is insensitive to various model specifications, mitigating the concern that certain model hyperparameters are driving the results. We also train a one-dimensional CNN model with firms' time-series of raw earnings data and document that the out-of-sample drift-predicting performance is inferior to that of our two-dimensional CNN model, thus emphasizing the importance of image representation.

Our study make several contributions to the literature. First, we contribute to a growing literature studying the applications of machine learning techniques in asset pricing. In particular, machine learning can efficiently combine information in a large set of characteristics to predict cross-sectional stock returns (Rapach et al., 2013; Kelly et al., 2019; Feng et al., 2020; Freyberger et al., 2020; Kozak et al., 2020; Gu et al., 2020; Gu et al., 2021; Leippold et al., 2022; Cao et al., 2024; Chen et al., 2024; Murray et al., 2024), mutual fund alphas (DeMiguel et al., 2023; Kaniel et al., 2023), hedge fund returns (Wu et al., 2021), bond risk premiums (Bianchi et al., 2021; Kelly et al., 2023), and option returns (Büchner and Kelly,

 $^{^{3}}$ We use the standardized unexpected earnings (SUE) to calculate the earnings growth measure.

⁴Prior studies employing the Mishkin test framework include Sloan (1996), Dechow and Sloan (1997), Rangan and Sloan (1998), Collins and Hribar (2000), Narayanamoorthy (2006), Cao and Narayanamoorthy (2012), Chen and Shane (2014), Hui et al. (2016), Ma and Markov (2017), and He and Narayanamoorthy (2020).

2022; Bali et al., 2023). A key differentiating feature of our CNN approach from the above studies is that we do not require an expansive list of input variables for training. Instead, our input is a two-dimensional image plotted using the most recent eight quarterly earnings, and the out-of-sample drift-predicting performance is still significant.⁵

We particularly contribute to a burgeoning literature employing CNN to extract relevant information from images. For example, Obaid and Pukthuanthong (2022) extract information from a large sample of news media images and translate that information into a daily investor sentiment index. Jiang et al. (2023) extract return-predicting information from stock-level charts depicting daily open, close, high, low prices, as well as trading volume and average prices over a past period to forecast future returns. Motivated by the two studies, we utilize CNN to examine whether certain features in earnings images are predictive of post-earnings announcement drift, thus also contributing to a vast literature studying the relation between earnings-related characteristics and post-earnings announcement performance. However, a notable distinction between our paper and these studies (e.g., Ball and Brown, 1968; Foster et al., 1984; Bernard and Thomas, 1989; Chan et al., 1996; Rangan and Sloan, 1998; Francis et al., 2004; Jegadeesh and Livnat, 2006; Livnat and Mendenhall, 2006; Narayanamoorthy, 2006; Kishore et al., 2008; Cao and Narayanamoorthy, 2012; Akbas et al., 2017; Kausar, 2018; He and Narayanamoorthy, 2020) lies in the approach to locate the earnings-related characteristics. In particular, we do not resort to subjective feature-engineering, but instead rely on CNN to *automatically* detect features that are most indicative of post-earnings announcement drift from earnings images in the training phase.

⁵Transforming data from one-dimensional to two-dimensional can potentially create more nuanced information, thus adding flexibility in prediction tasks when the input data is scarce.

Our paper also sheds light on a recent literature examining the information value in visualized data. For example, Nekrasov et al. (2022) find that visuals in firms' Twitter earnings announcements are associated with more retweets, representing increased attention to the earnings news. Moss (2022) find that retail investors use their visual perception of earnings surprise displayed on Robinhood rather than the unexpected earnings scaled by stock price in their investment decisions. Hu and Ma (2023) quantify persuasion in visual, vocal, and verbal dimensions in start-up pitch videos, and find that passionate and warm pitches significantly increase funding probability. Cao et al. (2024) examine the visual information in corporate executive presentations and examine how market participants respond to such information. Christensen et al. (2024) document a significant increase in the disclosure of infographics in 10-K filings over time, and investigate the relation between the use of infographics and uncertainty in capital markets. Gu et al. (2024) find that a daily firm-level investor sentiment measure based on graphics interchange format images (GIFs) in postings about firms on Stocktwits.com is positively correlated with same-day stock returns while predicting stock return reversals in the following two weeks. Our paper, on the other hand, proposes a universal approach to visualize a firm's time-series of quarterly earnings into a bar-chart image that accounts for both the sign and magnitude of earnings as an input to machines. To the best of our knowledge, we are the first to systematically visualize earnings data and extract information from earnings images to predict post-earnings announcement drift.

The rest of the paper is structured as follows. Section 2 describes the data and variables. Section 3 describes how we generate earnings images, assign labels, and train the CNN model. Section 4 presents the out-of-sample drift-predicting performance of the CNN model. Section 5 examines the nature of CNN predictions and the source of their drift-predicting power. Section 6 performs robustness checks. Finally, Section 7 concludes the paper.

2 Data and Variables

We focus on U.S. common stocks traded on NYSE, AMEX, and NASDAQ, and obtain data from Compustat and CRSP. First, we collect Compustat firm-quarters whose earnings announcement date (Compustat item RDQ) is between January 1974 and June 2023, and delete observations with missing RDQ in the most recent eight quarters (quarters q-7 to q). Next, we apply filters in He and Narayanamoorthy (2020) to eliminate announcements that are potentially subject to data errors. In particular, we delete observations if in the most recent eight quarters, a firm has (i) more than one earnings announcement on any date (ii) earnings announcement date within 30 days of a previous earnings announcement date, or (iii) earnings announcement either prior to or more than 180 days after the corresponding fiscal period-end.

We require a firm to have non-missing earnings in the most recent eight quarters and a CRSP daily price higher than one dollar at the most recent earnings announcement date (quarter q). We use income before extraordinary items (Compustat item IBQ) as earnings. Financial and utility firms with SIC codes from 6000 to 6999 and from 4900 to 4949 are excluded. In addition, firms are required to have non-missing market capitalization (SIZE) and non-negative book-to-market ratio (BM), and have at least 90 non-missing daily return observations in the [-120, -31] window relative to the current quarter earnings announcement date. We are left with 404,635 firm-quarter observations after applying all the above filters.

Next, we define the in-sample dataset and the out-of-sample dataset. The in-sample dataset consists of 124,413 firm-quarter observations between January 1974 to December 1993 (4,548 firms; 80 quarters), and is for CNN model training and validation. The complement out-of-sample dataset is for predicting and testing the out-of-sample CNN model performance, and thus serves as the dataset for all empirical analyses throughout the paper. It consists of 240,844 firm-quarter observations between July 1994 to December 2023 with non-missing firm characteristics (7,527 firms; 116 quarters).⁶

We summarize the definitions of all the variables used in this study, in the Appendix. To mitigate the impact of outliers, we transform most variables into decile ranks (numbered 0 to 9, from low to high) following prior research (e.g., Rangan and Sloan, 1998; Livnat and Mendenhall, 2006; Garfinkel and Sokobin, 2006).⁷ The cutoff points for quarterly variables are based on the distribution of those quarterly variables in the previous quarter. The cutoff points for annual variables from July in year t to June in year t + 1 are based on the distribution of those annual variables at the end of June in year t. Then, we convert all the decile ranks to scaled ranks by dividing by 9 and subtracting 0.5. The resulting scaled ranks vary from -0.5 to 0.5 with a mean of zero and a range of one. This variable transformation approach is to facilitate comparison of the economic magnitudes of firm characteristics. For

⁶The firm characteristics include SUE, EA, TREND, RET[-1, 1], RET[-30, -2], PERSIST, VOL, GP, OP, OA, TA, and AG. Along with BM and SIZE, these firm characteristics are used as comparing/control variables throughout the paper. See Appendix for variable definitions. In addition, we address the reasons to set up a six-month lag between the end of the in-sample dataset and the out-of-sample dataset in Section 4.1. In addition, our empirical results are robust to using an in-sample (out-of-sample) period consisting of 60 quarters (136 quarters) or 100 quarters (96 quarters).

⁷Variables that are not transformed into decile ranks are the six measure of the 63-day post-announcement buy-and-hold abnormal return (BHAR), including market-adjusted return (MAR), size-adjusted return (SAR), and four factor-adjusted returns (FF4, FF6, HMXZ5, and DHS3).

example, the coefficient on a variable of interest (in scaled rank) in a return regression represents the return from a zero investment strategy of going long on the highest variable decile and short on the lowest variable decile.

3 The CNN Model

In this section, we introduce the CNN training procedure, which can be regarded as an image classification task. First, we transform firms' times series of quarterly earnings into bar charts. Then, we assign labels to each earnings bar-chart image based on the relative performance of its post-earnings announcement drift among the cross-section of firms in the same quarter. Finally, we train the CNN model with the 124,413 earnings images in the in-sample period (1974Q1 to 1993Q4, 80 quarters) to "learn" drift-predicting information.

3.1 Generating earnings images

We begin by plotting the most recent eight quarterly earnings in bar charts. Following Jiang et al. (2023), we generate black-and-white rather than colored images for simplicity and uniformity. Each black-and-white image is of size 24×24 pixels, which is recognized by the machine as a 24×24 matrix of 0 (black pixel) and 255 (white pixel). We use black as the background color and white as the color for earnings, and the constant image size setup is for better comparison of earnings patterns across different firms in different quarters.

Each quarter occupies 24×3 pixels in the image, and quarterly earnings are plotted as "white bars" in the middle column of each quarter. In particular, let $E_1, E_2, ..., E_8$ denote the most recent eight quarterly earnings corresponding to quarter $q-7, q-6, ..., q, E_{\text{MAX}}$ and E_{MIN} denote the maximum and minimum of the eight quarterly earnings, and r() denote the function that rounds the input value to the nearest whole number. We set the bottom-left vertex of the image as the origin of a two-dimensional coordinate system, so a rectangular area in the image can be represented as $([x_1, x_2], [y_1, y_2])$. Next, we classify firms' most recent eight quarterly earnings into one of the three types, determine the values corresponding to the top and bottom of the image, and plot each quarterly earnings into bars accordingly. The three types are as follows:

• Type I ($E_{\text{MIN}} \ge 0$; the most recent eight quarterly earnings are all non-negative): In this case, we set E_{MAX} and 0 as the top and bottom of the image, respectively. E_i is plotted as the area of

$$\left(\left[3i-2,3i-1\right],\left[0,r(24*\frac{E_i}{E_{\text{MAX}}})\right]\right),\tag{1}$$

for i = 1, ..., 8. Figure 1 displays an example earnings image of this type. The maximum earnings is E_7 and thus it occupies a whole column. All other quarterly earnings are plotted upward, and their heights are determined using E_7 as the reference point.

• Type II ($E_{\text{MAX}} > 0$ and $E_{\text{MIN}} < 0$; the maximum quarterly earnings is positive while the minimum earnings is negative): In this case, E_{MAX} and E_{MIN} coincide with the top and bottom of the image, respectively. The implicit "zero-earnings line" corresponds to $r(24 * \frac{-E_{MIN}}{E_{MAX} - E_{MIN}})$, and E_i is plotted above or below the zero-earnings line as follows:

$$\begin{cases} \left[3i-2,3i-1\right], \left[r(24*\frac{-E_{MIN}}{E_{MAX}-E_{MIN}}), r(24*\frac{-E_{MIN}}{E_{MAX}-E_{MIN}}) + r(24*\frac{E_i}{E_{MAX}-E_{MIN}})\right] & \text{if } E_i > 0, \\ \left[3i-2,3i-1\right], \left[r(24*\frac{-E_{MIN}}{E_{MAX}-E_{MIN}}) - r(24*\frac{-E_i}{E_{MAX}-E_{MIN}}), r(24*\frac{-E_{MIN}}{E_{MAX}-E_{MIN}})\right] & \text{if } E_i \le 0, \\ (2) \end{cases}$$

for i = 1, ..., 8. Figure 2 displays an example earnings image of this type. Here we see the advantage of using bar charts as opposed to line graphs when plotting earnings. Bars can represent positive, zero, or negative earnings without further specifying numbers on the vertical axis. Positive earnings are plotted upward while negative earnings are plotted downward, and the bar lengths (in pixels) are computed as the rounded value of 24 multiplied by the absolute values of E_i scaled by $E_{\text{MAX}} - E_{\text{MIN}}$.

Type III (E_{MAX} ≤ 0; the most recent eight quarterly earnings are all non-positive): In this case, 0 and E_{MIN} coincide with the top and bottom of the image, respectively. E_i is plotted as the area of

$$\left[3i - 2, 3i - 1\right], \left[24 - r(24 * \frac{E_i}{E_{\text{MIN}}}), 24\right],$$
(3)

for i = 1, ..., 8. Figure 3 displays an example earnings image of this type. The minimum earnings is E_5 and thus it occupies a whole column. All earnings are plotted downward, and their heights are plotted using E_5 as the reference point.

Note that in all three types, it is possible for E_i to be very close to zero after scaling and thus does not occupy a full pixel in the image after rounding, i.e., $y_1 = y_2$.⁸ In addition,

 $^{^{8}}$ One extreme case is that all eight quarterly earnings are very close to each other so that when plotting

the distance between two neighboring earning of pixel between is greater than the distance between the leftmost (or rightmost) earnings and the border of the image, which is consistent with the default setup of a bar chart for most statistical software.

3.2 Assigning labels to earnings images

Next, we assign one of the three labels ("sell", "hold", or "buy") to each firm's earnings image based on the firm's 63-day post-announcement buy-and-hold abnormal return (BHAR). There are different ways to define buy-and-hold abnormal returns, and we use market-adjusted buy-and-hold returns (MAR).⁹ In particular, $MAR_{i,q+1}$ is defined as the difference between the buy-and-hold return of firm i and that of the CRSP value-weighted market portfolio over the windows [2, 64] in trading days relative to firm i's earnings announcement date t in quarter q:

$$MAR_{i,q+1} = \prod_{\tau=2}^{64} (1 + R_{i,t+\tau}) - \prod_{\tau=2}^{64} (1 + R_{M,t+\tau}), \qquad (4)$$

where R_i is the delisting-adjusted return of firm i, R_M is the return of the CRSP value-weighted market return, and t is quarter q's announcement date of firm i.¹⁰ The 63-day holding window corresponds to the total number of trading days in three months. We follow previous studies (Vega, 2006; Engelberg et al., 2012; Frank and Sanati, 2018) to compute MAR from day 2 to mitigate the impact of bid-ask bounce and other market microstructure effects, and our results are robust to MAR defined using the trading window of [1, 63].

earnings on a bar chart, each earnings bar occupies a whole column. In this case, one cannot tell from the image whether all earnings are positive or negative. However, we checked all earnings images and did not find this extreme case.

⁹In untabulated tests, we find that the final CNN out-of-sample performance is robust to using alternative definitions of abnormal returns such as size-adjusted or factor-adjusted returns in the label-assigning process.

¹⁰We replace missing delisting-adjusted returns with market returns, which is equivalent to reinvesting any remaining proceeds in the market portfolio until the end of the holding period.

Then, for each quarter in the in-sample period (1974Q1 to 1993Q4, 80 quarters), we sort firms announcing earnings into terciles based on their 63-day MAR. The bottom, mid, and top terciles are labeled "sell", "hold", and "buy", respectively. Since the number of training images for each label is about the same, we mitigate the class imbalance issue in CNN training that arises with a disproportionate ratio of labels.

3.3 CNN architecture and training

In this section, we introduce the general algorithm of CNN and describe the architecture and training process of our CNN model.

3.3.1 CNN architecture

A CNN model typically consists of multiple building blocks, with each building block consisting of a convolutional layer and a pooling layer.¹¹ In the convolutional layer, an input image is first scanned by a set of convolutional filters to generate feature maps. Convolutional filters, also known as kernels, are small matrices (usually of size 3×3 , 5×5 , or 7×7 pixels) that are applied over the input image's pixels to detect features. The matrix elements in convolutional filters are also known as "weights", which will later be optimized during the training process of the CNN model.

Then, a filter "scans" an image. It starts at the top-left corner of an image and moves one pixel at a time. In CNN terminology, this corresponds to a "stride" of 1, which is usually the default option in convolution. In each position, an element-wise multiplication is performed between the filter and the corresponding patch of the input image. The products

¹¹A building block may also consist of multiple convolutional layers, working sequentially (as in our case) or in parallel (e.g., GoogLeNet (Szegedy et al., 2015)).

are summed to a single value placed in the corresponding position of the feature map.¹² The process is repeated until the filter slides across the entire image, thus generating a complete feature map. Then, the feature map passes through an activation function to introduce non-linearity.¹³ Finally, in the pooling layer, pooling operations are applied to the feature map produced by the convolutional layer (after activation) to retain the most important information.

Figure 4 illustrates the operation of a CNN building block using a black-and-white 6×6 pixels input image and a 3×3 pixels convolutional filter as an example. We first apply "padding" to the input image by filling the absent neighbor elements with zeros for elements at the image's border, which helps preserve input dimensions and allow better edge feature detection.¹⁴ Then, the convolutional layer applies the filter to the input image and produces a feature map of size 6×6 pixels. In particular, we see that the top-left (bottom-right) 3×3 pixels patch of the input image, after applying the 3×3 pixels convolutional filter, eventually becomes the top-left (bottom-right) element in the output feature map.

Next, we use "Leaky ReLU" as our activation function, which is a variation of the conventional ReLU function.¹⁵ In particular, Leaky ReLU function transforms an input value x to itself if x > 0 and 0.01x otherwise. We see that in Figure 4, Leaky ReLU function is applied to every element of the feature map generated by the convolutional layer, and

¹²In CNN terminology, a feature map is the output produced after applying a convolutional filter to an input image or the previous layer's feature map.

¹³Without activation functions, CNNs would just consist of linear operations (matrix multiplication).

¹⁴Convolution operations without padding inevitably reduce the spatial dimensions of the output feature maps.

¹⁵Compared to ReLU, Leaky ReLU allows a small, non-zero gradient for negative input values, which helps to address the "dying ReLU" problem (where some neurons can become permanently inactive during training) and thus enables more robust learning in neural networks. Interested readers may refer to Maas et al. (2013), which is the first modern deep learning reference to Leaky ReLU.

a feature map of the same size $(6 \times 6 \text{ pixels})$ is produced. Finally, we employ the most commonly used 2×2 pixels max-pooling filter with a stride of 2 as our pooling operation function. The max-pooling filter scans the feature map produced by the convolutional layer (after activation), selects the maximum element in the 2×2 pooling window, and eventually shrinks the height and width of the input feature map by half. Hence, max-pooling helps preserve the most prominent features and reduce the spatial dimensions of the input feature maps.

Having introduced the components of a CNN building block, we proceed to describe the architecture of our CNN model and illustrate the details in Figure 5. In particular, we use three building blocks, with the first block consisting of 64 convolutional filters of 7×7 pixels, the second block consisting of 128 convolutional filters of 3×3 pixels, and the third block consisting of 256 convolutional filters of 3×3 pixels.¹⁶ Since our input images are black-and-white with low resolution (24×24 pixels), we do not resort to an overly complicated CNN model with too many blocks. In addition, we employ filters of 7×7 pixels in the first block to ensure that pixels of neighboring quarterly earnings in the initial input earnings image can always be scanned simultaneously by the convolutional filters.¹⁷

After passing an image sequentially through the three building blocks, the CNN model "flattens" the elements in the feature maps generated by the last block to a vector.¹⁸ Then,

¹⁶We follow the literature and increase the number of filters after each convolutional layer by a factor of two. Following several well-known CNN architectures (e.g., VGGNet (Simonyan and Zisserman, 2014); ResNet (He et al., 2016); DenseNet (Huang et al., 2017)), we employ 64 filters in the first convolution layer. The choice of 64 filters provides a balance between model complexity and computational efficiency, making it a popular choice.

¹⁷This filter size choice is based on the hypothesis that certain patterns in neighboring quarterly earnings are helpful in predicting post-earnings announcement drift, although in Table 9 we find that the main results are robust to using filters of 3×3 or 5×5 pixels in the first block.

¹⁸In the second and third building blocks where the input is a "stack" of feature maps instead of a single image, the output will be a stack of feature maps as well. In particular, each filter is applied to the stack of feature maps to perform convolution and eventually generate one feature map. Hence, the total number of

the fully connected layer linearly transforms the elements in the vector to three scores $(Z_1,$ Z_2 , and Z_3) of the three labels (label 1 = "sell", label 2 = "hold", label 3 = "buy").¹⁹ Finally, the Softmax function transforms these scores to three probabilities $(\hat{y}_1, \hat{y}_2, \hat{y}_3)$, where $0 \leq \hat{y}_i \leq 1$ and $\sum_{i=1}^3 \hat{y}_i = 1.^{20}$ Hence, one can interpret \hat{y}_3 as the CNN-predicted likelihood of an earnings image to be classified as "buy" when the other two labels are available. Throughout the paper, we refer to \hat{y}_3 as the *CNN buy* probability (CNNBP).

3.3.2**CNN** training

CNN training is about finding the optimized weights, i.e., parameters in convolutional filters and the fully connected layer, to minimize model "loss" to a certain extent. We follow the CNN literature to use the cross-entropy loss function as the loss function for minimization. In particular, let $y = [y_1, y_2, y_3]'$ denote the label of an earnings image, which is either [1, 0, 0]', [0, 1, 0]', or [0, 0, 1]' corresponding to the sell, hold, and buy label, respectively. The cross-entropy loss is computed as

$$Loss(y, \hat{y}) = -\sum_{i=1}^{3} y_i * \log \,\hat{y}_i,$$
(5)

where loss $\in [0, \infty)$ and smaller loss represents better CNN performance.

We closely follow the regularization procedures in Gu et al. (2020) and Jiang et al. (2023) to train our CNN model.²¹ From the in-sample period (1974Q1 to 1993Q4, 80 quarters), we

feature maps in the output is equal to the number of filters. In CNN terminology, the number of stacks is usually referred to as the "depth" of the input.

¹⁹The linear transformation also requires parameters to be estimated and optimized in the training process.

²⁰The Softmax function converts the three scores into a probability distribution of three outcomes, i.e., $\hat{y}_i = \frac{e^{Z_i}}{\sum_{k=1}^3 e^{Z_k}}$, for i = 1, 2, 3. ²¹Interested readers may refer to Gu et al. (2020) for detailed explanations on those modeling choices.

In particular, we use batch normalization (Ioffe and Szegedy, 2015) to mitigate the internal covariate shift

randomly select 70% earnings images for training and the other 30% for validation, which are labeled training dataset and validation dataset, respectively. First, 128 earnings images (batch size = 128) are randomly selected from the training dataset and are passed through the CNN model as described in Figure 5 to produce the average loss. The loss is propagated back through the model to update the weights via stochastic gradient descent and the Adam algorithm (Kingma and Ba, 2014) with a learning rate of 10^{-5} . Next, the model randomly selects 128 images from the remaining images in the training dataset to update the weights. The iteration process stops when the model sees all earnings images in the training dataset. We then apply those updated weights to the earnings images in the validation dataset to compute validation loss. After completing the above process, we finish training an "epoch".

Next, we start the training process again using the updated weights of the first epoch as the initial weights, and eventually obtain updated weights and validation loss of the second epoch. This training iteration process is halted only when the validation loss fails to improve for two consecutive epochs, and the updated weights of the third-to-last epoch are stored as the optimized weights.²² Then, we can then apply these optimized weights to a new earnings image to generate CNN-predicted likelihood for a label of interest.

4 CNN Performance

In this section, we examine the out-of-sample performance of the CNN trained model. We begin by applying the CNN-trained weights to the earnings images in 1994Q2 to 2023Q2 to

problem, choose Xavier initializers (Glorot and Bengio, 2010) as initial weights for model training, and apply a 50% dropout rate (Srivastava et al., 2014) to the fully connected layer to prevent over-fitting. In Table 9, we show that none of these choices affect our main results in Table 2.

 $^{^{22}}$ This technique is called 'early stopping", which is to prevent over-fitting to the training data.

generate the *CNN buy* probability.²³ Since CNN training can result in different outcomes even when using the same architecture and dataset due to the stochastic nature of optimization algorithms and the use of dropout rate, we train the same CNN model independently for ten times (number of ensembles = 10) and then average the *CNN buy* probability, which helps in achieving better accuracy and robustness.

If the CNN model is capable of extracting features that are indicative of post-earnings announcement performance (*CNN buy* features) from earnings images, there should be a positive relation between the *CNN buy* probability and post-earnings announcement drift. Hence, we examine whether firms with higher *CNN buy* probability on average experience higher post-earnings announcement drift, and whether such return-predictive power shares resemblance with that of the existing return predictors.

4.1 Portfolio analysis: univariate sort

4.1.1 CNN buy probability and post-earnings announcement drift

In each quarter starting from 1994Q3, we assign firms announcing earnings into decile portfolios based on their *CNN buy* probability, where the cutoffs are based on the distribution of the previous quarter's *CNN buy* probability. This approach (Foster et al., 1984; Bernard and Thomas, 1989) is to prevent hindsight bias that classifies firms into portfolios based on information not available at the time when a strategy is implemented. Hence, the out-of-sample period is from 1994Q3 to 2023Q2 (a total of 116 quarters). Next, we compute the average 63-day MAR for each *CNN buy* probability decile. If the CNN model is

 $^{^{23}\}mathrm{We}$ start from 1994Q2 as late announcers in 1993Q4 require post-earnings announcement drift data in 1994Q1 to form labels.

competent in detecting *CNN buy* features from earnings images, the average 63-day MAR should be monotonically increasing when going from the lowest to highest *CNN buy* probability deciles.

Table 2 presents the results. We find that when moving from the lowest decile to the highest decile of CNN buy probability, the average CNN buy probability²⁴ increases from 25.5% to 44.4%, and the average 63-day MAR increases monotonically from -0.6% to 3.0%. The average difference in MAR is 3.6% (*t*-statistic = 7.213) in a quarter, which corresponds to an annualized return exceeding 14%. Figure 6 further depicts the return differential for each of the 116 quarters in the out-of-sample period. Specifically, the hedge return is positive in 99 out of the 116 quarters (85.3%), indicating that the CNN out-of-sample performance is stable over time.

We next study whether the return differential is robust to alternative risk adjustment other than the market model by examining the relation between the CNN buy probability and the 63-day size-adjusted buy-and-hold returns (SAR) as well as factor-adjusted buy-and-hold returns. In particular, SAR is defined as the difference between the buy-and-hold return of an announcing firm and that of a size-matched portfolio over the 63-day window ([2, 64]) following its earnings announcement date. We use the monthly NYSE size decile breakpoints at the end of June in year t to determine the size-matched portfolio for a firm whose earnings announcement date is between July of year t to June of year t + 1. Monthly size breakpoints and daily size portfolio returns are obtained from Kenneth French's website.

To compute the 63-day factor-adjusted buy-and-hold returns, we replace $R_{M,k}$ in equation (4) with daily returns $\widehat{R_{F,k}}$ predicted by factor models. To compute $\widehat{R_{F,k}}$, we first estimate

 $^{^{24}}$ As a benchmark, random label assignments would generate a CNN buy probability of 1/3.

individual stock factor loadings by regressing returns on the factors on a 120-day rolling window from t - 150 to t - 31 for each stock:

$$r_{i,t} = \alpha_i + \beta_i' F_t + \epsilon_{i,t},\tag{6}$$

where $r_{i,t}$ is the excess return on stock *i* and F_t is a vector of factors. The predicted return $\widehat{R_{F,k}}$ is then computed as $\widehat{\beta}'_i F_k$.²⁵ In particular, we consider the factors in the Fama-French four- and six-factor models (Fama and French, 1993; Carhart, 1997; Fama and French, 2015; Fama and French, 2018), the q^5 -model (Hou et al., 2015; Hou et al., 2021), and the risk-and-behavioral model (Daniel et al., 2020). The 63-day factor-adjusted buy-and-hold returns following an earnings announcement of these models are denoted FF4, FF6, HMXZ5, and DHS3, respectively.²⁶

Columns 4 to 10 in Table 2 present qualitatively similar results when we compute the average 63-day SAR or factor-adjusted buy-and-hold returns (FF4, FF6, HMXZ5, and DHS3) for each CNN buy probability decile. In particular, the return differential between the highest and lowest CNN buy probability deciles range from 3.1% to 3.5%, with t-statistics all statistically significant at the 1% level. Overall, Table 2 shows a significantly positive relation between CNN buy probability and post-announcement buy-and hold abnormal returns that

²⁵See, for example, Savor (2012) and Kapadia and Zekhnini (2019).

²⁶Fama and French (2015) extends the Fama-French three-factor model (Fama and French, 1993) to control for operating profitability (RMW) and investment (CMA). After the inclusion of a momentum factor (Carhart, 1997), we have Fama-French four-factor and six-factor models (Fama and French, 2018). Hou et al. (2015) propose the q-model to control for market, size (ME), investment (IVA), and profitability (return on equity, ROE), and Hou et al. (2021) further includes an expected growth factor (EG) into the q^5 -model. Daniel et al. (2020) propose a 3-factor risk-and-behavioral model that accounts for market, long-term financing (FIN), and short-term earnings surprise (PEAD). Fama-French factors are obtained from Kenneth French's website, q^5 -model factors are obtained from Lu Zhang's website, and DHS3 factors are obtained from Lin Sun's website.

are robust to various risk adjustments.²⁷

4.1.2 Firm characteristics and post-earnings announcement drift

Next, we examine whether the *CNN buy* probability is superior to the other usual determinants of drift. We first consider three earning attributes: standardized unexpected earnings (Ball and Brown, 1968; Bernard and Thomas, 1989; Foster et al., 1984), earnings acceleration (He and Narayanamoorthy, 2020), and trend in gross profitability (Akbas et al., 2017). In particular, standardized unexpected earnings (SUE) reflects earnings surprise based on a seasonal random walk model, earnings acceleration (EA) captures the change in earnings growth from one quarter to the next, and trend in gross profitability (TREND) characterizes the recent path in a firm's profitability in addition to the profit level.

In addition to the three earnings attributes, we also consider market capitalization (Fama and French, 1992, 1993), book-to-market ratio (Fama and French, 1992, 1993), earnings announcement return (Foster et al., 1984; Chan et al., 1996), pre-announcement return (Carhart, 1997), earnings persistence (Francis et al., 2004), earnings volatility (Cao and Narayanamoorthy, 2012), gross profitability (Novy-Marx, 2013), operating profitability (Ball et al., 2016), total accruals (Richardson et al., 2005), operating accruals (Sloan, 1996; Hribar and Collins, 2002), and asset growth (Cooper et al., 2008).

We follow the approach in the previous section to assign firms announcing earnings into decile portfolios based on one of the 14 firm characteristics, where the cutoffs are based on the distribution of the previous quarter's firm characteristic. Then, we examine the difference

 $^{^{27}}$ The results are qualitatively the same if we assign earnings images with labels based on firms' 21-day or 42-day post-announcement market-adjusted buy-and-hold returns, train the CNN model, and examine the 21-day or 42-day post-announcement buy-and-hold abnormal returns for CNN buy probability deciles.

in the average 63-day MAR, SAR, and factor-adjusted buy-and-hold returns (FF4, FF6, HMXZ5, and DHS3) between the highest and lowest deciles of each characteristic. If a firm characteristic is more successful in predicting post-earnings announcement drift, the difference in the post-announcement buy-and-hold abnormal returns between the highest and lowest deciles sorted on the characteristic should be larger in magnitude.

Table 3 reports the results. The first row presents the return differential between the highest and lowest deciles sorted on CNN buy probability (ranging from 3.1% to 3.6%), which is the same as the last row in Table 2. In comparison, we find that the statistically significant return differential ranges from 2.0% to 2.6% for SUE, from 2.1% to 2.4% for EA, from 1.4% to 2.0% for TREND, from 3.3% to 3.8% for RET[-1, 1], from 1.6% to 2.5% for BM, from 1.0% to 2.4% for GP, and from -1.9% to -1,6% for AG. For the other firm characteristics, the return differential fails to remain statistically significant at the 10% level across all six abnormal return measures. Overall, Table 3 suggests that the drift-predicting power of the CNN buy features is superior to that of the usual determinants of PEAD, while being roughly on par with that of the earnings announcement return (RET[-1, 1]).

4.2 Portfolio analysis: double sorts

In this section, we examine whether the drift-predicting power of CNN buy features is distinct from that of the 14 firm characteristics. In particular, we construct 5×5 portfolios sorted independently (Liu et al., 2018; He and Narayanamoorthy, 2020) on the CNN buy probability and one of the six firm characteristics (SUE, EA, TREND, RET[-1, 1], BM, GP, and AG). Again, to alleviate hindsight bias, we use the distribution of each firm characteristic in the previous quarter to form the quintile cutoffs. Then, we examine the average 63-day MAR of the 25 portfolios.²⁸

We present the double sorts results in Table 4. Consistent with previous findings in the literature, in Panel A we find that firms with high SUE outperform firms with low SUE by a quarterly return of 0.9% to 2.5% depending on the *CNN buy* probability quintile. On the other hand, the return differential between the high and low *CNN buy* probability quintiles is significantly positive across medium to high SUE quintile while being insignificantly positive in the bottom two quintiles. This finding indicates that the *CNN buy* features exhibit incremental drift-predicting power for medium to high SUE firms, while this predictive ability appears to be subsumed by the SUE effect for low SUE firms.

In Panel B we present the analogous two-way sorting based on firms' earnings acceleration. We find that the average difference in 63-day MAR between the top and bottom CNN buy probability quintiles ranges from 1.6% to 5.4% (t-statistics between 2.711 and 5.954) across all EA quintiles. In Panel C, we find similar evidence: the positive relation between the CNN buy probability and 63-day MAR is not limited to any TREND quintiles. In particular, the hedge return based on the CNN buy probability appears to be larger in magnitude in the low and high quintiles of EA and TREND. We also find that the CNN buy features subsume some return predictability of EA and TREND. Turning to Panels D, E, F, and G we find that the average difference in 63-day MAR between the highest and lowest CNN buy probability quintiles remains positive and statistically significant at the 1% level across all RET[-1, 1], BM, GP, and AG quintiles.

We report the double sorts results based on the other eight characteristics in Table IA1 of

²⁸The results are robust to alternative return measures (SAR, FF4, FF6, HMXZ5, and DHS3).

the Internet Appendix, and find that the drift-predicting power of *CNN buy* features persists in all the quintiles of the other seven characteristics. The results in Table 4 and Table IA1 combined indicate that CNN's ability to predict post-earnings announcement drift is mostly distinct from that of the usual determinants.

4.3 Cross-sectional regression

We next perform a cross-sectional regression analysis to simultaneously control for the firm characteristics that may affect the positive relation between the *CNN buy* probability and post-earnings announcement drift. Following prior literature (Akbas, 2016), we estimate quarterly weighted Fama and MacBeth (1973) regressions in which the dependent variable is the firm's 63-day MAR. We begin by first running the following cross-sectional regression every quarter:

$$MAR_{i,q+1} = \alpha_q + \beta_q CNN \ buy \ probability_{i,q} + \sum \beta_{c,q} Controls_{i,q} + \varepsilon_{i,q+1}, \tag{7}$$

where *i* refers to the stock, *q* refers to the quarter, and the *CNN buy* probability and control variables are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of zero. Then, we average the cross-sectional coefficients across all quarters, where the weights correspond to the number of observations in each quarterly cross-sectional regression. In addition to using MAR as the dependent variable, we also employ SAR and four factor-adjusted buy-and-hold returns (FF4, FF6, HMXZ5, and DHS3).

Table 5 presents the regression results. The coefficient on the CNN buy probability in Column 1 is 0.014 (t-statistic = 3.529), suggesting that a long-short strategy of going long on the highest *CNN buy* probability decile and short on the lowest decile generates a 63-day MAR of around 1.4%, controlling for other firm characteristics simultaneously. In columns 2 to 5 where we replace MAR with SAR and factor-adjusted buy-and-hold returns, the coefficients on the *CNN buy* probability range from 0.010 to 0.014 and are all statistically significant at the 1% level, suggesting that our results are not caused by omission of risk factors.

In addition, we find that in all model specifications, the post-earnings announcement drift is significantly increasing in earnings announcement return (RET[-1, 1]) and operating profitability (OP) while decreasing in market capitalization (SIZE).²⁹ Overall, the results in Table 5 provide strong support for the out-of-sample drift-predicting power of the *CNN buy* features, which cannot be accounted for by stock anomalies or lack of risk controls.

5 Interpreting CNN Predictions

So far, we've demonstrated that *CNN buy* features possess significant drift-predicting power. In this section, we first explore the nature of these *CNN buy* features via linear approximation. Then, we examine whether *CNN buy* features exhibit incremental predictive ability for future earnings growth, and whether the drift-predicting power of *CNN buy* features can be attributable to market investors missing such predictive ability.

²⁹While SUE and RET[-1,1] both proxy for earnings surprises, their coefficients remain positively significant, consistent with Kishore et al. (2008)'s findings that trading strategies formed based on SUE and RET[-1,1] are largely independent of each other.

5.1 Linear approximation of CNN predictions

We begin by linearly fitting CNN predictions with firm characteristics and historical earnings. In particular, in Table 6 we estimate quarterly weighted Fama and MacBeth (1973) regressions in the out-of-sample period (1994Q3-2023Q2, 116 quarters) using *CNN buy* probability as the dependent variable. The independent variables are the 14 firm characteristics considered before in specification 1 and earnings in the most recent eight quarters in specification 2.³⁰

In specification 1, we find that CNN buy probability is positively correlated with standardized unexpected earnings (SUE), earnings acceleration (EA), earnings announcement return (RET[-1,1]), pre-announcement return (RET[-30, -2]), earnings volatility (VOL), book-to-market ratio (BM), gross profitability (GP), and asset growth (AG), while negatively related to earnings persistence (PERSIST), market capitalization (SIZE), operating profitability (OP), and total accruals (TA). The result suggests the CNN model is capable of discerning some meaningful return-predicting information, such as the earnings surprise effect, the SUE effect, the earnings acceleration effect, the gross profitability effect, the value effect, and the size effect solely from historical earnings represented in the form of images. However, the $CNN \ buy$ probability appears to load negative (positively) on operating profitability (asset growth) despite the fact that it has been documented to positively (negatively) predict subsequent returns.

Economically, the coefficient on SUE (= 0.510) is the largest and exceeds the second-largest coefficient (= -0.274) on SIZE by around a half. In particular, moving from the lowest decile to the highest decile of SUE (SIZE) is associated with a 51.0% (27.4%) incremental

 $^{^{30}}$ We follow Ball et al. (2009) to use return on assets (ROA) as the earnings measure, where ROA is defined as the quarterly earnings scaled by total assets in the previous quarter.

increase (decrease) in the decile of *CNN buy* probability. Overall, the 14 firm characteristics collectively explain 32.8% of the variation in *CNN buy* probability, indicating most of the variation in *CNN buy* features is left unexplained.

Turning to specification 2, we find that the $CNN \ buy$ probability is most associated with the current quarter's ROA and the ROA four quarters prior in terms of economic magnitude. In particular, moving from the lowest decile to the highest decile of ROA in the current quarter (four quarters prior) leads to a 64.4% (42.8%) incremental increase (decrease) in the decile of $CNN \ buy$ probability. This result to some extent explain why the $CNN \ buy$ features share the most resemblance with SUE among all the considered firm characteristics in specification 1. The eight historical earnings collectively explain 30% of the variation in $CNN \ buy$ probability, implying that 70% of the variation in $CNN \ buy$ probability is attributable to the nonlinear transformation of the underlying historical earnings.

In order to provide more insights on the reasoning behind CNN predictions, Figure 7 displays earnings images whose *CNN buy* probabilities rank in the top and bottom 15 among the 240,844 earnings images in the out-of-sample period (1994Q3 to 2023 Q2, 116 quarters). We find that earnings images with the highest *CNN buy* probabilities all have the current earnings as the maximum earnings and mostly have the earnings four quarters prior as the minimum earnings. In addition, there seems to be an increasing trend of quarterly earnings. In contrast, we see that earnings images with the lowest *CNN buy* probabilities mostly have the current earnings as the minimum earnings images with the lowest *CNN buy* probabilities mostly have the current earnings as the minimum earnings. And their earnings in the previous two to four quarters are relatively high.

5.2 CNN predictions and future earnings growth

In this section, we examine whether the *CNN buy* features possess incremental predictive ability for future earnings growth and whether the post-announcement abnormal return based on *CNN buy* probability is associated with this predictive ability. Since the earnings images in Section 3 are plotted in a way similar to those investors would see during earnings call conferences or generate on their own via statistical software, we conjecture that the return-predicting power of *CNN buy* features likely manifests since market investors do not fully incorporate the implications of *CNN buy* features contained in earnings images for future earnings growth.

First, we run a regression of one-quarter-ahead earnings growth on *CNN buy* probability (CNNBP) in the out-of-sample period (1994Q3-2023Q2, 116 quarters). We use the previously defined SUE as the earnings growth measure (Bernard and Thomas, 1989; Ball and Bartov, 1996). In particular, we have

$$SUE_{q+1} = \alpha + \gamma_1 CNN \ buy \ probability_q + \sum \gamma_c Controls_q + \delta_{q+1},$$
 (8)

where SUE_{q+1} represents one-quarter-ahead earnings growth. Control variables are the 14 firm characteristics considered before.³¹ In columns 1 and 2 of Table 7, the coefficients on CNNBP are positive and statistically significant at the 1% level, suggesting that *CNN buy* probability is a significant predictor of earnings growth in the subsequent quarter. Economically, moving from the lowest decile to the highest decile of CNNBP in the current

 $^{^{31}}$ In particular, SUE_q serves as the control for the well-documented earnings autocorrelation pattern in prior studies.

quarter leads to a 13.3% incremental increase in the decile of one-quarter-ahead earnings growth, controlling for past earnings growth and other firm characteristics. On the other hand, the coefficients on SUE are significantly positive, consistent with the previous findings in the literature.

We next examine whether the *CNN buy* probability (CNNBP) helps predict the three-day abnormal return around the one-quarter-ahead earnings announcement $(\text{RET}[-1, 1]_{q+1})$.³² If this is the case, then investors do not appear to incorporate fully the implications of earnings acceleration for earnings announcement. We find that in columns 3 and 4 of Table 7, the coefficients on CNNBP are positive and highly significant, indicating a positive relation between *CNN buy* probability and the three-day abnormal return around the next earnings announcement date. In terms of economic magnitude, moving from the lowest decile to the highest decile of CNNBP in the current quarter leads to a 0.4% incremental increase in $\text{RET}[-1, 1]_{q+1}$.

Since *CNN buy* features have positive implications for both one-quarter-ahead earnings growth and for the three-day abnormal return surrounding the next earnings announcement date, we employ an econometrics framework, i.e., the Mishkin test (Mishkin, 1983; Abel and Mishkin, 1983) widely used in the earnings-based anomaly literature, to test whether the market fully understands the implications of the *CNN buy* features for SUE_{q+1} .³³ In particular, we simultaneously estimate two equations: an earnings forecasting equation and a rational pricing equation. In our context, the earnings forecasting equation is equation (8)

³²Shorter-window returns are typically less susceptible to risk considerations (Bernard and Thomas, 1990; Sloan, 1996; Narayanamoorthy, 2006; Cao and Narayanamoorthy, 2012)

³³See, for example, Sloan (1996), Dechow and Sloan (1997), Rangan and Sloan (1998), Collins and Hribar (2000), Narayanamoorthy (2006), Cao and Narayanamoorthy (2012), Chen and Shane (2014), Hui et al. (2016), Ma and Markov (2017), and He and Narayanamoorthy (2020).

that characterizes the evolution of earnings growth.

Next, for the rational pricing equation, we assume a linear abnormal return (AR) model (e.g., Sloan, 1996) that satisfies the efficient-markets condition:

$$AR_{q+1} = \beta(SUE_{q+1} - SUE_{q+1}^e) + \varepsilon_{q+1}, \tag{9}$$

where β is a multiple, $\text{SUE}_{q+1}^e = E_q(\text{SUE}_{q+1})$ is the rational forecast of SUE_{q+1} in quarter q, and ε_{q+1} is a noise in quarter q+1 satisfying $E_q(\varepsilon_{q+1}) = 0$. In equation (9), abnormal returns are zero in expectation, i.e., $E_q(\text{AR}_{q+1}) = 0$, and market efficiency implies that only $(\text{SUE}_{q+1} - \text{SUE}_{q+1}^e)$, the unanticipated changes in SUE, can be correlated with AR_{q+1} . In other words, if the market correctly understands the implications of the *CNN buy* features for future earnings growth as depicted in equation (8), AR_{q+1} should only be related to the earnings growth surprise $(\text{SUE}_{q+1} - \text{SUE}_{q+1}^e)$, but not related to the *CNN buy* probability in quarter q.

Combining the earnings growth forecasting model in equation (8) with the rational pricing model in equation (9) provides the following system:

Forecasting equation:
$$SUE_{q+1} = \alpha + \gamma_1 CNNBP_q + \sum \gamma_c Controls_q + \delta_{q+1}$$
 (10)

Pricing equation:
$$\operatorname{AR}_{q+1} = \beta(\operatorname{SUE}_{q+1} - \alpha^* - \gamma_1^* \operatorname{CNNBP}_q - \sum \gamma_c^* \operatorname{Controls}_q) + \varepsilon_{q+1}.$$
 (11)

The two systems are simultaneously estimated using iterative-weighted non-linear least squares (Mishkin, 1983), and the coefficients with * represent the coefficients inferred from

market investors' expectation of SUE_{q+1} .³⁴ In particular, we are interested in testing whether $\gamma_1 = \gamma_1^*$ holds or not, i.e., whether the observed relation between SUE_{q+1} and CNN buy features is the same as the relation between SUE_{q+1} and CNN buy features implicit in AR_{q+1} . In other words, $\gamma_1 = \gamma_1^*$ indicates that investors are fully aware of the implications of CNN buy features for SUE_{q+1} , and this restriction yields a likelihood ratio test statistic that has a chi-square distribution with one degree of freedom.³⁵ If $\gamma_1 = \gamma_1^*$ is rejected while $0 < \gamma_1^* < \gamma_1$, then investors only partially incorporate the implications of CNN buy features for future earnings growth. On the other hand, if $\gamma_1 = \gamma_1^*$ is rejected and $\gamma_1^* = 0$, then investors appear to completely ignore the implications of CNN buy features for future earnings growth.

We report the estimated coefficients, t-statistics based on firm and quarter double-clustered standard errors, and likelihood ratio test statistics of the Mishkin test in Table 8. In particular, AR_{q+1} is either the abnormal return from a three-day window around quarter q + 1's earnings or the quarter-long window starting two days after the quarter q earnings and ending on the next announcement date. All variables except for the abnormal return AR are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of 0.

Since $\gamma_1 = \gamma_1^*$ is rejected at the 1% level (likelihood ratio statistic 8.218 for the three-day window and 21.444 for the quarter-long window) and since $\gamma_1 > \gamma_1^*$, it appears that market

 $^{^{34}}$ Kraft et al. (2007) shows that the exclusion of control variables from the forecasting and pricing equations leads to an omitted variables problem. That is, if the variables omitted are not rationally priced and are also correlated with the variable of interest in the forecasting equation, then the source of market inefficiency cannot be correctly identified. Hence, we include various control variables that may be related to *CNN buy* probability.

³⁵The test statistic of the Mishkin test is $2 \times n \times ln(SSR^c/SSR^u)$ distributed asymptotically $\chi^2(q)$, where q is the number of constraints imposed by market efficiency, n is the number of observations in each equation, SSR^c is the sum of squared residuals from the constrained weighted system, and SSR^u is the sum of squared residuals from the unconstrained weighted system.

investors are underestimating the implications of CNN buy features for future earnings growth. In particular, the quarter-long window γ_1^* (= 0.010) is statistically indistinguishable from zero³⁶ while the three-day window γ_1^* (= 0.062) is highly significant, implying that market investors are completely ignoring the positive implications of CNN buy features for future earnings growth at the time of the current earnings announcement, but partially understands these implications by the time of the next earnings announcement. In other words, the market gradually learns more about the implications of CNN buy features for future earnings growth from other sources of information by the time of the next earnings announcement.

Overall, the results in this section provide evidence that the positive relation between the *CNN buy* probability and post-earnings announcement drift is consistent with market investors not fully understanding the implications of the *CNN buy* features for one-quarter-ahead earnings growth.

6 Robustness Checks

6.1 Month-based rebalancing trading strategy

The out-of-sample tests in Table 2 involves buying and selling stocks two days after an earnings announcement, which requires significant attention and thus may be difficult to implement in reality. Hence, in this section we examine whether sorting stocks based on the

³⁶In untabulated tests we use SAR_{q+1} and four factor-adjusted buy-and-hold returns (FF4_{q+1}, FF6_{q+1}, HMXZ5_{q+1}, and DHS3_{q+1}) to measure quarter-long AR_{q+1}, respectively. We find that in all specifications, $\gamma_1 = \gamma_1^*$ is rejected at the 1% level, and γ_1 is statistically indistinguishable from zero. Hence, the fact that market investors are completely unaware of the implications of CNN buy features can not be attributed to lack of risk controls.

CNN buy probability to form a more conservative month-based rebalancing trading strategy (Hou et al., 2020; Jensen et al., 2023) can still generate profits.

In particular, at the end of each month t in the out-of-sample period, we sort firms into deciles based on the *CNN buy* probability computed using the most recent eight quarterly earnings. For a firm to enter the portfolio formation at the end of month t, we require that announcement date of the most recent earnings to be within three months prior to portfolio formation to exclude stale earnings information. We then examine the average returns in the subsequent month t + 1 for each *CNN buy* probability decile.

Table 9 presents the equal-weighted and value-weighted average portfolio returns for each *CNN buy* probability decile. Panel A shows that a hedge portfolio going long in the top CNN-based buy probability decile and short in the bottom decile yields an average equal-weighted monthly return of 1.0%. The factor-adjusted hedge returns range from 0.7% to 0.8% and are all statistically significant at the 1% level. On the other hand, the average value-weighted hedge returns are significantly positive but are smaller in magnitude (ranging from 0.3% to 0.5%). This is because CNN model is treating each input image equally during the training phase, regardless of market capitalization. Hence, CNN predictions are unsurprisingly more accurate when we employ CNN stored parameters to form out-of-sample portfolios with equal-weights rather than value-weights.

6.2 Alternative CNN modeling choices

We next explore whether the main results in Table 2 are sensitive to model specifications. In particular, we re-train the CNN model with alternative modeling choices, as listed in the first column of Table 10, and then examine the return differential in the 63-day post-announcement BHAR between the highest and lowest *CNN buy* probability deciles in the out-of-sample period.

In Panel A, we experiment with different combinations of filter size and number of convolution layers. The combination of our CNN model can be expressed as $(7 \times 7, 3 \times 3, 3 \times 3)$, while we consider alternative modeling choices of $(5 \times 5, 3 \times 3, 3 \times 3)$, $(5 \times 5, 3 \times 3, 3 \times 3)$, $(7 \times 3, 3 \times 3)$, $(5 \times 5, 3 \times 3)$, and $(3 \times 3, 3 \times 3)$. We find that the 63-day MAR, SAR, and factor-adjusted returns remain positive and highly significant, indicating that CNN model performance is mostly insensitive to those choices. In addition, omitting the batch normalization step or Xavier initialization, adjusting the activation function from leaky ReLU to ReLU, or lowering the dropout rate from 0.5 to 0 does not generate a noticeable loss in performance either. Hence, our main results are robust to alternative modeling choices.

In Panel B, we employ a one-dimensional CNN model in training where the inputs are 1×8 pixels row vectors consisting of the time-series of firms' most recent eight quarterly earnings (in the form of ROA) as inputs, and the convolutional filters sliding across the inputs are row vectors as well.³⁷ In other words, the one-dimensional CNN model is a special case of the two-dimensional CNN model, with both the inputs and the convolutional filters shrinking from matrices to row vectors. In particular, we consider the following modeling choices: $(1 \times 7, 1 \times 3, 1 \times 3), (1 \times 5, 1 \times 3, 1 \times 3), (1 \times 3, 1 \times 3), (1 \times 5, 1 \times 3, 1 \times 3), (1 \times 3, 1 \times 3), (1 \times 7, 1 \times 3), (1 \times 5, 1 \times 3), (1 \times 5, 1 \times 3, 1 \times 3), (1$

³⁷The results are robust to using unscaled earnings numbers.

than one-third of that in Panel A. The results suggest that image representation of historical earnings produces more information useful in prediction post-earnings announcement drift.

7 Conclusion

In this study, we examine the drift-predicting information contained in visualized earnings data. In particular, we apply CNN to earnings images plotted using time series of quarterly earnings to *automatically* extract *CNN buy* features that are most predictive of post-earnings announcement drift. In out-of-sample tests, we find that firms in the highest *CNN buy* probability decile significantly outperform firms in the lowest *CNN buy* probability decile by 3.6% in the 63-day post-announcement window. In addition, the drift-predicting power of *CNN buy* features is robust to a battery of controls for risk, distinct from that of the previously documented anomalies and earnings attributes, and stable over time.

We find that while *CNN buy* probability shares some resemblance with firm characteristics known to predict returns, its variation is largely left unexplained. In particular, the *CNN buy* probability appear to be positively associated with one-quarter-ahead earnings growth as well as the three-day abnormal return surrounding the next earnings announcement. As a result, we employ a direct market efficiency test and find that high abnormal returns following high *CNN buy* probability and the positive implications of *CNN buy* features for future earnings growth are strongly associated. In other words, the drift-predicting power of *CNN buy* features is consistent with investors not incorporating fully the implications of *CNN buy* features for future earnings growth.

In addition, the drift-predicting ability of CNN buy features persists in a more conservative

monthly-rebalancing strategy setting, and remains insensitive to various model specifications when image representation is used. Overall, our paper highlights the usefulness of applying deep learning techniques to visualized data in predicting post-earnings announcement returns. **Appendix. Variable Definitions.** This table summarizes variable definitions. Compustat annual or quarterly items are colored in blue.

Variables	Descriptions
MAR	Market-adjusted return (MAR) is defined as the difference between the buy-and-hold return of an announcing firm and that of the CRSP value-weighted market portfolio over the 63-day windows [2,64] following its earning announcement date.
SAR	Size-adjusted return (SAR) is defined as the difference between the buy-and-hole return of an announcing firm and that of a size-matched portfolio over the 63-day window ([2,64]) following its earnings announcement date. We use the monthly NYSE size decile breakpoints at the end of June in year t to determine the size-matched portfolio for a firm whose earnings announcement date is between July of year t to June of year $t + 1$.
FF4	Fama-French three-factor and momentum-adjusted buy-and-hold return during the 63-day window ([2,64]) following earnings announcement date, with facto loadings estimated using the 120-day window ([-150,-31], 90 days minimum) prio to the earnings announcement date. The factors are market, size, value, and momentum.
FF6	Fama-French five-factor and momentum-adjusted buy-and-hold return during the 63-day window ([2,64]) following earnings announcement date, with factor loadings estimated using the 120-day window ([-150,-31], 90 days minimum) prior to the earnings announcement date. The factors are market, size, value, operating profitability, investment, and momentum.
HMXZ5	q5-factor-adjusted buy-and-hold return during the 63-day window ([2,64] following earnings announcement date, with factor loadings estimated using the 120-day window ([-150,-31], 90 days minimum) prior to the earning announcement date. The factors are market, size, investment, return on equity and expected growth.
DHS3	Behavioral-factor-adjusted buy-and-hold return during the 63-day window ([2,64] following earnings announcement date, with factor loadings estimated using the 120-day window ([-150,-31], 90 days minimum) prior to the earning announcement date. The factors are market, financing, and post earning announcement drift.
SUE	Standardized unexpected earnings, defined as the change in split-adjusted quarterly earnings per share $\left(\frac{\text{EPSPXQ}}{\text{AJEXQ}}\right)$ from its value four quarters ago divided by the standard deviation of this change over the prior eight quarters (six quarter minimum). SUE also serves as the earnings growth proxy.

Variables	Descriptions					
EA	Earnings acceleration. For firm i in quarter q , we use					
	$\frac{\text{EPS}_{i,q} - \text{EPS}_{i,q-4}}{\text{Stock Price}_{i,q-1}} - \frac{\text{EPS}_{i,q-1} - \text{EPS}_{i,q-5}}{\text{Stock Price}_{i,q-2}},$					
	where $\text{EPS}_{i,q}$ is earnings per share for firm <i>i</i> in quarter <i>q</i> . Shares are adjusted for stock splits.					
TREND	Trend in quarterly gross profitability. For firm i in quarter q , we use $\beta_{i,q}$ estimate from the following time-series regression:					
	$GPQ_{i,q} = \alpha_{i,q} + \beta_{i,q}t + \lambda_{1,i,q}D1 + \lambda_{2,i,q}D2 + \lambda_{3,i,q}D3 + \epsilon_{i,q},$					
	where $t = 1, 2,, 8$ and represents a deterministic time trend covering quarter q -through q , and D1 to D3 represent quarterly dummy variables. GPQ is calculate as sales revenue (SALEQ) minus costs of goods sold (COGSQ), divided by tota assets (ATQ). If SALEQ is unavailable, we use quarterly revenue (REVTQ) If COGSQ is unavailable, we use quarterly total operating expenses (XOPRQ minus quarterly selling, general and administrative expenses (XSGAQ, zero missing).					
$\operatorname{RET}[-1,1]$	Earnings announcement return, defined as the value-weighted market-adjuste stock return during the $[-1, 1]$ window around earnings announcement date.					
RET[-30, 2]	Pre-announcement return, defined as the value-weighted market-adjusted stock return during the $[-30, -2]$ window prior to earnings announcement date.					
PERSIST	Earnings persistence. For firm <i>i</i> in quarter <i>q</i> , we use $\beta_{i,q}$ estimated from th following time-series regression:					
	$\text{EARNINGS}_{i,q} = \alpha_{i,q} + \beta_{i,q} \text{EARNINGS}_{i,q-1} + \epsilon_{i,q},$					
	with the most recent eight quarters (quarter $q - 7$ to q) of earnings (IBQ).					
VOL	Earnings volatility. We use the standard deviation of earnings (IBQ) in the most recent eight quarters (quarter $q - 7$ to q).					
SIZE	Firm size for July of year t to June of year $t + 1$ is defined as June market capitalization (from CRSP) of year t .					
BM	Book-to-market ratio for July of year t to June of year $t + 1$ is defined as boo equity for the fiscal year ending in calendar year $t - 1$ divided by the market capitalization at the end of December of $t - 1$. Book equity is computed a stockholders' book equity (SEQ), plus deferred taxes (TXDB, zero if missing and investment tax credit (ITCB, zero if missing), minus the book valu of preferred stock (depending on availability, we use redemption (PSTKRF) carrying (PSTKL), or par value (PSTK)).					

Variables	Descriptions
GP	Gross profitability for July of year t to June of year $t+1$ is defined as sales revenue (SALE) minus cost of goods sold (COGS), divided by total assets (AT) for the fiscal year ending in calendar year $t-1$. If SALE is unavailable, we use revenue (REVT). If COGS is unavailable, we use total operating expenses (XOPR) minus selling, general and administrative expenses (XSGA, zero if missing).
OP	Operating profitability for July of year t to June of year $t + 1$ is defined as sales revenue (SALE) minus cost of goods sold (COGS), minus selling, general, and administrative expenses (XSGA), and plus research and development expenditures (XRD, zero if missing), scaled by total assets (AT) for the fiscal year ending in calendar year $t - 2$. If SALE is unavailable, we use revenue (REVT). If COGS is unavailable, we use total operating expenses (XOPR) minus selling, general and administrative expenses (XSGA, zero if missing).
ТА	Total accruals for July of year t to June of year $t + 1$ is defined as net income (NI) minus operating, investing, and financing net cash flows (OANCF, IVNCF, and FINCF) plus sales of stocks (SSTK, zero if missing) minus stock repurchases and dividends (items PRSTKC and DV, zero if missing) for the fiscal year ending in calendar year $t - 1$, scaled by total assets (AT) for the fiscal year ending in $t - 2$.
OA	Operating accruals for July of year t to June of year $t+1$ is defined as net income (NI) minus net cash flow from operations (OANCF) for the fiscal year ending in calendar year $t-1$, scaled by total assets (AT) for the fiscal year ending in $t-2$.
AG	Asset growth for July of year t to June of year $t+1$ is defined as total assets (AT) for the fiscal year ending in calendar year $t-1$ minus total assets for the fiscal year ending in $t-2$, scaled by total assets for the fiscal year ending in $t-2$.
ROA	Return on assets is defined as quarterly earnings (IBQ) divided by total assets (ATQ) in the previous quarter.

	E_1	E_2	E_3	E_4	E_5	E_6	E_7	E_8
Earnings (in millions) $\begin{bmatrix} x_1, x_2 \end{bmatrix}$ $\begin{bmatrix} y_1, y_2 \end{bmatrix}$	$3.163 \ [1,2] \ [0,10]$	$\begin{array}{c} 4.882 \\ [4,5] \\ [0,16] \end{array}$	5.068 [7,8] [0,17]	$\begin{array}{c} 4.472 \\ [10,11] \\ [0,15] \end{array}$	$\begin{array}{c} 4.243 \\ [13,14] \\ [0,14] \end{array}$	5.263 [16, 17] [0, 17]	$7.348 \\ [19, 20] \\ [0, 24]$	$\begin{array}{c} 6.542 \\ [22,23] \\ [0,21] \end{array}$

Figure 1. An Example of Type I's Earnings Image. This figure displays a black-and-white 24×24 pixels earnings image for a firm whose quarterly earnings in the most recent eight quarters (quarters q - 7 to q) are all non-negative. $E_1, E_2, ...,$ and E_8 represent the quarterly earnings in quarter q - 7, q - 6, ..., and q, respectively. The bottom-left vertex of an image is set as the origin of a two-dimensional coordinate system, and a rectangular area in an image is represented as $([x_1, x_2], [y_1, y_2])$.

	E_1	E_2	E_3	E_4	E_5	E_6	E_7	E_8
Earnings (in millions) $\begin{bmatrix} x_1, x_2 \end{bmatrix} \\ \begin{bmatrix} y_1, y_2 \end{bmatrix}$	$-0.263 \\ [1,2] \\ [4,8]$	-0.609 [4,5] [0,8]	-0.110 [7,8] [6,8]	$0.114 \\ [10, 11] \\ [8, 10]$	$\begin{array}{c} 0.322 \\ [13,14] \\ [8,12] \end{array}$	$ \begin{array}{c} 1.122\\[16,17]\\[8,24]\end{array} $	$0.989 \\ [19, 20] \\ [8, 22]$	$\begin{array}{c} 0.945 \\ [22, 23] \\ [8, 21] \end{array}$

Figure 2. An Example of Type II's Earnings Image. This figure displays a black-and-white 24×24 pixels earnings image for a firm whose maximum quarterly earnings in the most recent eight quarters (quarters q - 7 to q) is positive, and the minimum quarterly earnings in the most recent eight quarters is negative. $E_1, E_2, ...,$ and E_8 represent the quarterly earnings in quarter q - 7, q - 6, ..., and q, respectively. The bottom-left vertex of an image is set as the origin of a two-dimensional coordinate system, and a rectangular area in an image is represented as $([x_1, x_2], [y_1, y_2])$.

	E_1	E_2	E_3	E_4	E_5	E_6	E_7	E_8
Earnings (in millions) $\begin{bmatrix} x_1, x_2 \\ y_1, y_2 \end{bmatrix}$	$-0.364 \\ [1,2] \\ [23,24]$	$-1.214 \\ [4,5] \\ [21,24]$	-1.763 [7,8] [20,24]	$-0.920 \\ [10, 11] \\ [22, 24]$	$-10.361 \\ [13, 14] \\ [0, 24]$	$-1.985 \\ [16, 17] \\ [19, 24]$	$-3.551 \\ [19, 20] \\ [16, 24]$	$\begin{array}{c} -4.016 \\ [22,23] \\ [15,24] \end{array}$

Figure 3. An Example of Type III's Earnings Image. This figure displays a black-and-white 24×24 pixels earnings image for a firm whose quarterly earnings in the most recent eight quarters (quarters q - 7 to q) are all non-positive. $E_1, E_2, ...,$ and E_8 represent the quarterly earnings in quarter q - 7, q - 6, ..., and q, respectively. The bottom-left vertex of an image is set as the origin of a two-dimensional coordinate system, and a rectangular area in an image is represented as $([x_1, x_2], [y_1, y_2])$.

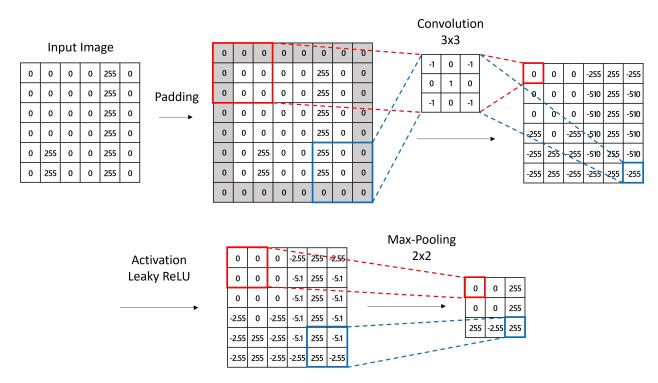


Figure 4. CNN Building Block: Padding, Convolution, Activation, and Max-Pooling. This figure displays how padding, convolution, activation, and max-pooling work in a CNN building block. The example input image is black-and-white and of size 6×6 pixels. The convolutional filter is of size 3×3 pixels. By using padding, the output after convolution has the same size of 6×6 pixels. The activation function is Leaky ReLU, which transforms an input value x to itself if x > 0 and 0.01x otherwise. The max-pooling (2×2 pixels) operation shrinks the input width and height to half by extracting the maximum element within a 2×2 pixels area and sliding through the image with a stride of 2.

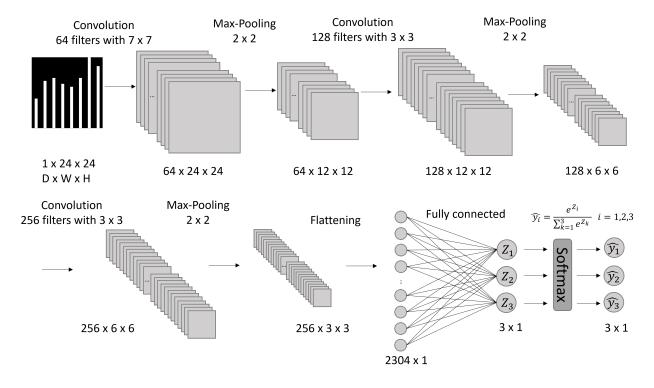
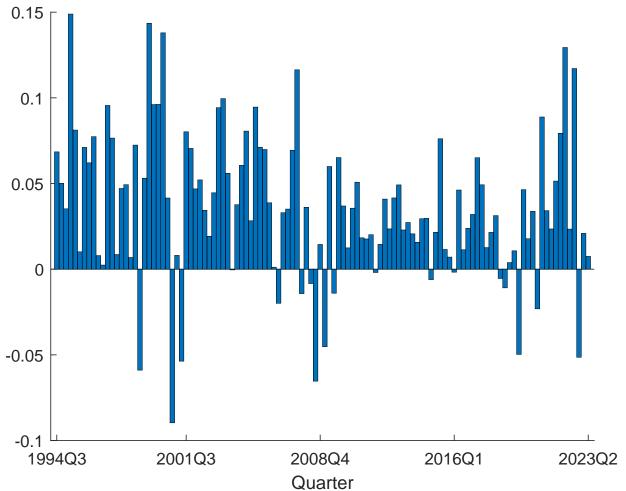


Figure 5. CNN Architecture Diagram. This figure displays the architecture of the CNN model. The notation $D \times W \times H$ represents the size of an image/feature map, where D is the depth, W is the width, and H is the height. The input black-and-white image is of size 24×24 pixels. There are three "blocks" in the model, with each block consisting of a convolutional layer and a max-pooling layer. The first convolutional layer has 64 filters of size 7×7 pixels, the second convolutional layer has 128 filters of size 3×3 pixels, and the third convolutional layer has 256 filters of size 3×3 pixels. After convolution, the output has the same width and height as those of the input due to padding, while its depth increases to the number of filters in the convolutional layer. After max-pooling, the output has half the width and height of the input, while its depth is the same as that of the input. Flattening refers to the process of converting the elements in a series of matrices into a vector. The fully connected layer linearly transforms the values in the vector to produce three "scores" of the three labels (label 1 = "sell", label 2 = "hold", label 3 = "buy"). Finally, the Softmax function transforms the three scores to three probabilities $(\hat{y}_1, \hat{y}_2, \hat{y}_3)$ that sum to one, and \hat{y}_3 is the *CNN buy* probability.



Difference in buy-and-hold MAR between high and low CNNBP deciles

Figure 6. Time Stability of CNN Out-Of-Sample Drift-Predicting Performance. This figure depicts the difference in the average 63-day buy-and-hold market-adjusted returns (MAR) between high and low *CNN buy* probability (CNNBP) deciles in each quarter during the out-of-sample period (1994Q3 to 2023Q2, 116 quarters). The decile cutoffs are based on the distribution of the previous quarter's *CNN buy* probability.

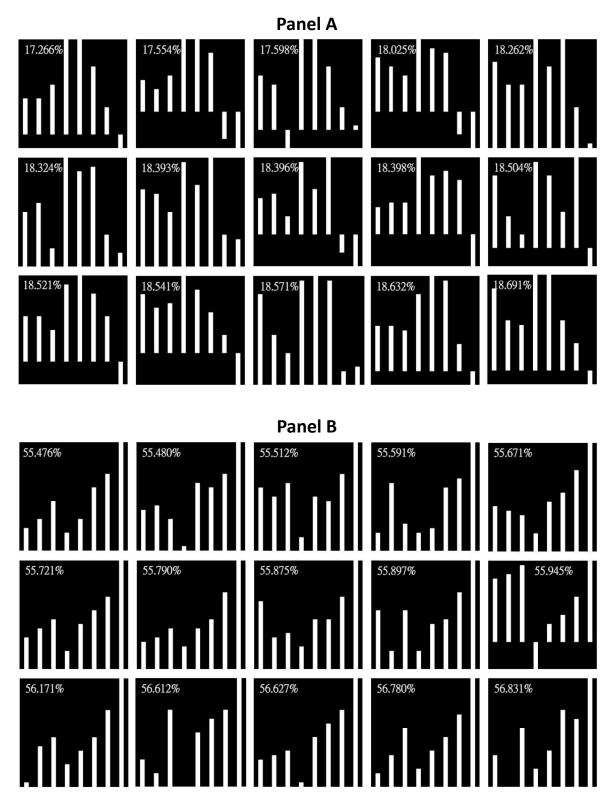


Figure 7. Earnings Images of the Lowest and Highest *CNN Buy* Probabilities. Panels A and B present earnings images whose *CNN buy* probabilities rank in the top 15 and bottom 15 among those of all earnings images in the out-of-sample period (1994Q3 to 2023Q2, 116 quarters), respectively. The corresponding *CNN buy* probabilities are also reported in each earnings image.

training and validation. The out-of-sample dataset is for testing the out-of-sample CNN model performance, and thus serves as the dataset for all empirical analyses throughout the paper. See Appendix for variable definitions.	thus serves
All Compustat firm-quarters with matched CRSP Permo (SHRCD = 10 or 11; EXCHCD = 1, 2 or 3) whose earnings announcement date (Compustat item RDQ) is between $1974/01/01$ and $2023/06/30$	815,907
Drop observations with missing RDQ in the most recent eight quarters	(86, 724)
Drop observations with earnings announcements on the same date for the same firm in the most recent eight quarters	(3, 349)
Drop observations with RDQ less than 30 day away from the previous quarter RDQ in the most recent eight quarters	(17,028)
Drop observations with RDQ before or more than 180 days after the quarter fiscal period end date in the most recent eight quarters	(2,971)
Drop observations with missing earnings (Compustat item IBQ) in the most recent eight quarters	(74,655)
Drop observations whose CRSP daily price at the current quarter RDQ is missing or $\leq \$1$	(102,680)
Drop financial firms (SIC codes between 6000 and 6999) and utility firms (SIC codes between 4900 and 4949)	(109, 418)
Drop observations with non-positive BM or missing SIZE	(14, 423)
Drop observations with more than 30 missing CRSP daily returns in the 120-day window ([-150, -31]) prior to the current quarter RDQ	(24)
Total observations	404,635
In-sample dataset: observations whose current quarter RDQ is between $1974/01/01$ and $1993/12/31$	124,413
Out-of-sample dataset: observations whose current quarter RDQ is between 1994/07/01 and 2023/06/30 and have non-missing SUE, EA, TREND, RET[-1, 1], RET[-30, -2], PERSIST, VOL, GP, OP, OA, TA, and AG	240,844

Table 2. CNN Buy Probability and Post-Earnings Announcement Drift: Univariate Portfolio Analysis. This table reports the average 63-day buy-and-hold abnormal return (BHAR) after earnings announcements, including market-adjusted return (MAR), size-adjusted return (SAR), and factor-adjusted returns (FF4, FF6, HMXZ5, and DHS3) for portfolios formed based on CNN buy probability (CNNBP) deciles in the out-of-sample period (1994Q3-2023Q2, 116 quarters). The CNNBP decile cutoffs are based on the distribution of the previous quarter's CNNBP. The average CNNBP for each CNNBP decile is reported in brackets. See Appendix for variable definitions. Newey and West (1987) t-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

	The 63-day post-announcement buy-and-hold abnormal return							
CNNBP deciles	MAR	SAR	FF4	FF6	HMXZ5	DHS3		
Low [25.5%]	-0.006 (-1.137)	-0.015^{***} (-5.210)	-0.007^{***} (-3.364)	-0.008^{***} (-3.622)	-0.007^{**} (-2.458)	0.001 (0.293)		
2 [28.8%]	-0.003 (-0.624)	-0.012^{***} (-4.714)	(-0.003) (-1.385)	-0.002 (-1.201)	-0.001 (-0.584)	$0.004 \\ (0.919)$		
3 [30.6%]	$0.004 \\ (0.651)$	-0.006^{**} (-2.610)	$0.001 \\ (0.270)$	$0.002 \\ (0.512)$	0.006^{*} (1.979)	0.012^{**} (2.150)		
4 [31.9%]	$0.003 \\ (0.449)$	-0.006^{*} (-1.689)	-0.001 (-0.215)	$0.001 \\ (0.116)$	$0.004 \\ (1.104)$	0.013^{*} (1.833)		
5 [33.1%]	$0.006 \\ (0.820)$	-0.004 (-1.359)	$0.002 \\ (0.518)$	$0.003 \\ (0.680)$	0.007^{*} (1.929)	0.015^{**} (2.123)		
6 [34.3%]	0.011 (1.408)	0.001 (0.232)	$0.007 \\ (1.560)$	0.007^{*} (1.714)	0.010^{***} (2.823)	0.019^{***} (2.667)		
7 [35.5%]	0.014^{*} (1.729)	0.004 (1.204)	0.009^{*} (1.946)	0.009^{**} (2.085)	$\begin{array}{c} 0.013^{***} \\ (3.407) \end{array}$	0.021^{***} (2.792)		
8 [37.1%]	0.015^{**} (2.117)	0.005 (1.624)	0.011^{***} (2.901)	$\begin{array}{c} 0.011^{***} \\ (2.813) \end{array}$	$\begin{array}{c} 0.015^{***} \\ (3.754) \end{array}$	0.022^{***} (3.238)		
9 [39.4%]	0.024^{***} (3.850)	$\begin{array}{c} 0.014^{***} \\ (5.323) \end{array}$	$\begin{array}{c} 0.017^{***} \\ (5.371) \end{array}$	$\begin{array}{c} 0.016^{***} \\ (5.226) \end{array}$	$\begin{array}{c} 0.021^{***} \\ (6.388) \end{array}$	0.029^{***} (4.909)		
High [44.4%]	0.030^{***} (4.994)	$\begin{array}{c} 0.020^{***} \\ (5.093) \end{array}$	$\begin{array}{c} 0.024^{***} \\ (6.555) \end{array}$	$\begin{array}{c} 0.024^{***} \\ (6.830) \end{array}$	$\begin{array}{c} 0.027^{***} \\ (7.375) \end{array}$	0.036^{***} (6.039)		
High-Low [18.9%]	0.036^{***} (7.213)	$\begin{array}{c} 0.035^{***} \\ (7.111) \end{array}$	$\begin{array}{c} 0.031^{***} \\ (8.092) \end{array}$	$\begin{array}{c} 0.032^{***} \\ (9.035) \end{array}$	$\begin{array}{c} 0.034^{***} \\ (8.378) \end{array}$	$\begin{array}{c} 0.034^{***} \\ (8.192) \end{array}$		

Table 3. Firm Characteristics and Post-Earnings Announcement Drift: Univariate Portfolio Analysis. This table reports the average 63-day post-announcement buy-and-hold abnormal return (BHAR) difference between the highest and lowest variable deciles in the out-of-sample period (1994Q3-2023Q2, 116 quarters). We use market-adjusted return (MAR), size-adjusted return (SAR), and factor-adjusted returns (FF4, FF6, HMXZ5, and DHS3) as BHAR measures. The variables include the *CNN buy* probability (CNNBP), standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), market capitalization (SIZE), book-to-market ratio (BM), earnings announcement return (RET[-1, -1]), pre-announcement return (RET[-30, -2]), earnings persistence (PERSIST), earnings volatility (VOL), gross profitability (GP), operating profitability (OP), total accruals (TA), operating accruals (OA), and asset growth (AG). The variable decile cutoffs are based on the distribution of the previous quarter's variable. See Appendix for variable definitions. Newey and West (1987) *t*-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

The 63-day post-announcement buy-and-hold abnormal return difference between the highest and lowest variable deciles								
Variable	MAR	SAR	FF4	FF6	HMXZ5	DHS3		
CNNBP	0.036***	0.035***	0.031***	0.032***	0.034***	0.034***		
	(7.213)	(7.111)	(8.092)	(9.035)	(8.378)	(8.192)		
SUE	0.026***	0.026***	0.021***	0.021***	0.021***	0.020***		
	(5.025)	(5.142)	(5.558)	(6.161)	(5.559)	(4.691)		
EA	0.024***	0.024***	0.022***	0.023***	0.021***	0.024***		
	(6.969)	(6.839)	(6.877)	(6.827)	(6.044)	(7.489)		
TREND	0.020***	0.018***	0.014***	0.014***	0.016***	0.015***		
	(3.353)	(3.154)	(2.974)	(3.073)	(3.188)	(2.789)		
$\operatorname{RET}[-1,1]$	0.035***	0.035***	0.038***	0.037***	0.035***	0.033***		
	(5.551)	(5.967)	(7.711)	(7.813)	(6.499)	(5.660)		
RET[-30, -2]	-0.003	0.002	0.005	0.006	0.002	-0.004		
	(-0.291)	(0.268)	(0.812)	(1.056)	(0.217)	(-0.572)		
PERSIST	-0.003	-0.003	-0.004	-0.002	-0.007^{*}	-0.002		
	(-0.983)	(-0.876)	(-1.127)	(-0.536)	(-1.792)	(-0.459)		
VOL	-0.006	-0.003	-0.018^{**}	-0.021^{***}	-0.022^{***}	-0.021^{***}		
	(-0.769)	(-0.413)	(-2.406)	(-2.898)	(-2.770)	(-2.740)		
SIZE	-0.017	-0.012^{*}	-0.026^{**}	-0.027^{**}	-0.032^{***}	-0.034^{**}		
	(-1.376)	(-1.956)	(-2.041)	(-2.293)	(-2.828)	(-2.517)		
BM	0.019^{*}	0.016^{*}	0.022^{***}	0.017^{**}	0.025^{***}	0.020^{**}		
	(1.692)	(1.685)	(2.625)	(2.177)	(2.899)	(2.073)		
GP	0.020***	0.020***	0.024^{***}	0.019***	0.010^{*}	0.016^{***}		
	(2.910)	(3.222)	(4.011)	(3.091)	(1.677)	(2.842)		
OP	0.017	0.018^{**}	0.017	0.015	0.001	0.010		
	(1.518)	(2.352)	(1.615)	(1.528)	(0.100)	(0.895)		
OA	-0.005	-0.005	0.000	-0.002	-0.003	-0.004		
	(-0.770)	(-0.763)	(-0.064)	(-0.441)	(-0.762)	(-0.743)		
ТА	-0.012	-0.010	-0.006	-0.007	-0.014^{**}	-0.011		
	(-1.406)	(-1.502)	(-0.734)	(-1.013)	(-2.140)	(-1.307)		
AG	-0.019^{**}	-0.017^{**}	-0.018^{**}	-0.016^{**}	-0.018^{**}	-0.017^{*}		
	(-2.108)	(-2.236)	(-2.166)	(-2.047)	(-2.532)	(-1.936)		

Table 4. CNN Buy Probability and Post-Earnings Announcement Drift: Double Sorts. This table reports the average 63-day buy-and-hold market-adjusted returns (MAR) after earnings announcements for portfolios formed based on the CNN buy probability quintiles and one of the six firm characteristics including standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), earnings announcement return (RET[-1,1]), book-to-market ratio (BM), gross profitability (GP), and asset growth (AG) using independent two-way sorting in the out-of-sample period (1994Q3-2023Q2, 116 quarters). The quintile cutoffs are based on the distribution of these variables in the previous quarter. See Appendix for variable definitions. Newey and West (1987) *t*-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

Panel A: Two-way	sorting, contro	lling for stand	lardized unexpe	ected earnings ((SUE)	
		(CNN buy proba	bility quintiles		
SUE quintiles	Low	2	3	4	High	High-Low
Low	-0.007	0.000	-0.008	-0.004	0.007	0.013
2	(-1.295)	(0.025)	(-0.920)	(-0.385)	(0.541)	(1.208)
2	(-0.005) (-0.967)	-0.003 (-0.565)	0.006 (0.612)	0.001 (0.088)	(0.007) (0.587)	(0.012) (1.265)
3	-0.005	0.005	0.010	0.016**	0.021^{***}	0.025***
-	(-0.836)	(0.567)	(1.325)	(2.147)	(3.040)	(4.088)
4	0.002	0.011	0.015^{*}	0.024***	0.028***	0.027^{***}
TT: 1	(0.291)	(1.336)	(1.936)	(3.137)	(4.368)	(4.044)
High	0.009	0.009	0.017^{**}	0.021^{***}	0.029^{***}	(0.019^{***})
	(1.425)	(1.217)	(2.377)	(3.087)	(5.399)	(3.311)
High-Low	0.016^{***}	0.009	0.025^{***}	0.024^{***}	0.021^{*}	
	(3.972)	(1.280)	(4.269)	(3.478)	(1.748)	
Panel B: Two-way	sorting, contro	lling for earning	ngs acceleration	n (EA)		
		(CNN buy proba	bility quintiles		
EA quintiles	Low	2	3	4	High	High-Low
Low	-0.006	0.002	0.008	0.007	0.029***	0.035***
	(-0.634)	(0.176)	(0.726)	(0.594)	(2.921)	(5.031)
2	-0.006	0.000	0.001	0.009	0.021***	0.027^{***}
	(-1.121)	(0.044)	(0.220)	(1.406)	(3.465)	(5.954)
3	0.000	0.000	0.007	0.009**	0.016***	0.016***
4	(0.028)	(0.014)	(1.574)	(1.992)	(3.158)	(2.711)
4	0.001	0.000	0.008	(0.014^{**})	(0.020^{***})	(0.019^{***})
High	$(0.175) \\ -0.013$	$(-0.019) \\ 0.017$	$(1.440) \\ 0.016$	$(2.397) \\ 0.027^{***}$	$(3.998) \\ 0.041^{***}$	$(3.746) \\ 0.054^{***}$
Iligh	(-1.072)	(1.179)	(1.450)	(2.684)	(5.286)	(5.505)
TT: 1 T	· /	× ,	· · · ·	· /	· /	(0.000)
High-Low	-0.007	0.015^{**}	0.008^{*}	0.020^{***}	(2.012^{**})	
	(-0.744)	(2.025)	(1.709)	(4.476)	(2.021)	
Panel C: Two-way	sorting, contro	olling for trend	in gross profit	ability (TREN	D)	
		(CNN buy proba	bility quintiles		
TREND quintiles	Low	2	3	4	High	High-Low
Low	-0.009	-0.002	0.004	0.012	0.023***	0.032***
	(-1.470)	(-0.197)	(0.480)	(1.300)	(3.193)	(6.792)
2	-0.002	0.001	0.001	0.008	0.024***	0.025^{***}
	(-0.332)	(0.136)	(0.089)	(1.157)	(4.092)	(5.332)
3	-0.005	-0.001	0.005	0.016**	0.021^{***}	0.026***
	(-0.924)	(-0.262)	(0.963)	(2.017)	(3.389)	(5.801)
4	-0.001	0.007	0.011	(0.016^{**})	(0.025^{***})	0.027^{***}
Uigh	(-0.272)	(1.125)	(1.534)	(2.392)	(3.910)	(5.750)
High	(-0.002) (-0.352)	$0.014 \\ (1.157)$	0.018^{*} (1.684)	$(0.018)^{*}$ (1.940)	(0.037^{***}) (4.740)	0.039^{***} (6.632)
	,	· /	· /	· · · ·	· /	(0.052)
High-Low	0.007	0.016**	0.014^{**}	0.005	0.013**	
	(1.424)	(2.206)	(2.357)	(0.963)	(2.037)	

Table 4. CNN BuyProbability and Post-Earnings Announcement Drift: Double Sorts.(continued)

			CNN buy probal	bility quintiles		
$\operatorname{RET}[-1,1]$ quintiles	Low	2	3	4	High	High-Low
Low	-0.011 (-1.555)	-0.003 (-0.292)	-0.005 (-0.493)	0.003 (0.275)	0.017^{**} (2.008)	0.028^{***} (5.094)
2	(-0.004) (-0.852)	0.001 (0.090)	0.005 (0.677)	0.005 (0.814)	0.019^{***} (3.083)	0.023^{***} (4.832)
3	(-0.002) (-0.487)	0.003 (0.511)	(0.003) (0.542)	0.014^{**} (2.156)	0.019^{***} (3.681)	0.021^{***} (5.625)
4	(-0.003) (-0.574)	(0.011) (1.510)	0.012^{*} (1.757)	(2.374)	0.025^{***} (4.220)	0.028^{***} (6.207)
High	(0.001) (0.088)	(1.123) (1.123)	(2.966)	(3.030^{***})	0.044^{***} (5.706)	0.044^{***} (5.913)
High-Low	0.011^{**} (2.033)	0.012^{*} (1.861)	0.035^{***} (6.177)	0.027^{***} (3.847)	0.027^{***} (4.030)	

Panel E: Two-way sorting, controlling for book-to-market ratio (BM)

	CNN buy probability quintiles							
BM quintiles	Low	2	3	4	High	High-Low		
Low	-0.003	-0.004	0.000	0.007	0.017***	0.020***		
	(-0.660)	(-0.575)	(0.045)	(1.146)	(2.876)	(3.111)		
2	0.000	0.002	0.005	0.012**	0.019***	0.019***		
	(0.036)	(0.373)	(0.687)	(2.083)	(3.578)	(4.478)		
3	-0.003^{-1}	0.007	0.008	0.013^{*}	0.025***	0.027***		
	(-0.477)	(0.869)	(1.333)	(1.841)	(3.884)	(6.189)		
4	-0.004	0.006	0.011	0.014	0.029***	0.033***		
	(-0.639)	(0.679)	(1.323)	(1.548)	(4.276)	(7.875)		
High	-0.013*	0.010	0.017	0.020*	0.044***	0.057***		
0	(-1.787)	(0.863)	(1.468)	(1.796)	(4.274)	(7.456)		
High-Low	-0.010	0.014	0.017	0.013	0.027**			
~	(-1.407)	(1.326)	(1.632)	(1.322)	(2.532)			

Panel F: Two-way sorting, controlling for gross profitability (GP)

			CNN buy proba	bility quintiles		
GP quintiles	Low	2	3	4	High	High-Low
Low	-0.016^{**}	-0.006	-0.005	0.002	0.018**	0.034***
2	$(-2.215) \\ -0.007$	$(-0.504) \\ 0.007$	$(-0.555) \\ 0.012^*$	$(0.185) \\ 0.010$	$(2.184) \\ 0.025^{***}$	$(4.747) \\ 0.032^{***}$
3	$(-1.238) \\ -0.004$	$\begin{array}{c}(0.944)\\0.003\end{array}$	$(1.722) \\ 0.015^*$	$(1.476) \\ 0.014^{**}$	$(3.790) \\ 0.030^{***}$	$(7.520) \\ 0.035^{***}$
4	$(-0.853) \\ 0.005$	$(0.462) \\ 0.006$	$(1.931) \\ 0.014^*$	$(2.007) \\ 0.019^{***}$	$(4.698) \\ 0.028^{***}$	$(6.854) \\ 0.024^{***}$
High	$(0.921) \\ -0.002$	$(0.985) \\ 0.009$	(1.947) 0.010	(2.699) 0.026^{***}	$(4.563) \\ 0.036^{***}$	(4.837) 0.038^{***}
	(-0.299)	(1.219)	(1.387)	(2.765)	(5.585)	(8.569)
High-Low	0.015^{**} (2.490)	0.014^{*} (1.751)	0.016^{***} (2.725)	0.024^{***} (3.718)	0.019^{**} (2.340)	

Table 4. CNN BuyProbability and Post-Earnings Announcement Drift: Double Sorts.(continued)

Panel G: Two-	-way sorting,	controlling for	asset growth (A	AG)		
			CNN buy proba	ability quintile	ès	
AG quintiles	Low	2	3	4	High	High-Low
Low	-0.006	0.009	0.014	0.018	0.037***	0.042***
2	$(-0.689) \\ -0.003$	$\begin{array}{c}(0.830)\\0.003\end{array}$	$(1.316) \\ 0.013^*$	$(1.633) \\ 0.017^{**}$	$(4.198) \\ 0.033^{***}$	$(8.064) \\ 0.036^{***}$
3	$(-0.599) \\ -0.002$	$(0.429) \\ 0.006$	$(1.696) \\ 0.008$	$(2.204) \\ 0.016^{**}$	$(4.643) \\ 0.026^{***}$	(7.288) 0.028^{***}
-	(-0.491)	(1.089)	(1.305)	(2.566)	(3.877)	(6.338)
4	$0.000 \\ (-0.092)$	$0.002 \\ (0.324)$	0.011 (1.588)	0.013^{*} (1.855)	0.023^{***} (4.508)	0.024^{***} (5.149)
High	-0.010 (-1.478)	-0.003 (-0.323)	-0.005 (-0.644)	0.005 (0.670)	0.019^{***} (3.295)	0.029^{***} (5.409)
High-Low	$-0.004 \\ (-0.690)$	-0.012 (-1.600)	-0.019^{***} (-2.878)	-0.013^{*} (-1.796)	-0.018^{***} (-2.773)	

Table 5. CNN Buy Probability and Post-Earnings Announcement Drift: Regression Analysis. The table presents results of quarterly weighted Fama and MacBeth (1973) regressions in the out-of-sample period (1994Q3-2023Q2, 116 quarters) using the 63-day buy-and-hold abnormal return (BHAR) after earnings announcements, including market-adjusted return (MAR), size-adjusted return (SAR), and factor-adjusted returns (FF4, FF6, HMXZ5, and DHS3), as the dependent variables. The weights correspond to the number of observations used in each quarterly cross-sectional regression. The independent variables include the CNN buy probability (CNNBP), standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), market capitalization (SIZE), book-to-market ratio (BM), earnings announcement return ($\operatorname{RET}[-1, -1]$), pre-announcement return ($\operatorname{RET}[-30, -2]$), earnings persistence (PERSIST), earnings volatility (VOL), gross profitability (GP), operating profitability (OP), total accruals (TA), operating accruals (OA), and asset growth (AG). All variables except for six measures of BHAR (MAR, SAR, FF4, FF6, HMXZ5, and DHS3) are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of zero. See Appendix for variable definitions. Newey and West (1987) t-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

	(1)	(2)	(3)	(4)	(5)	(6)
	MAR	SAR	FF4	FF6	HMXZ5	DHS3
Intercept	0.011*	0.000	0.007**	0.007**	0.010***	0.019***
	(1.707)	(0.216)	(2.098)	(2.308)	(3.677)	(3.068)
CNNBP	0.014***	0.014***	0.010***	0.011***	0.012***	0.014***
	(3.529)	(3.535)	(3.507)	(3.652)	(3.971)	(4.344)
SUE	0.011***	0.012***	0.011***	0.010***	0.010***	0.009**
	(2.951)	(3.009)	(3.406)	(3.146)	(2.840)	(2.389)
EA	0.013***	0.013^{***}	0.012***	0.013***	0.013***	0.014***
	(4.892)	(4.713)	(4.681)	(5.133)	(4.833)	(5.417)
TREND	0.016***	0.016***	0.012***	0.013***	0.012***	0.013***
	(4.769)	(4.717)	(3.797)	(3.679)	(3.961)	(4.291)
$\operatorname{RET}[-1,1]$	0.021***	0.023***	0.024***	0.024***	0.023***	0.021***
	(5.734)	(6.520)	(7.126)	(6.861)	(7.008)	(5.788)
RET[-30, -2]	-0.011^{**}	$-0.006^{-0.006}$	-0.003^{-1}	-0.003^{-1}	$-0.006^{-0.006}$	-0.010^{**}
	(-2.113)	(-1.207)	(-0.745)	(-0.685)	(-1.106)	(-2.094)
PERSIST	0.000	0.000	-0.001	-0.001	-0.003	0.001
	(-0.155)	(-0.117)	(-0.594)	(-0.281)	(-1.409)	(0.385)
VOL	0.014	0.014	0.009	0.005	0.009	0.010
	(1.281)	(1.356)	(1.217)	(0.824)	(1.207)	(1.212)
SIZE	-0.028^{*}	-0.023^{**}	-0.035^{**}	-0.035^{**}	-0.040^{***}	-0.042^{**}
	(-1.697)	(-2.298)	(-2.216)	(-2.283)	(-2.789)	(-2.562)
BM	0.014	0.014	0.017***	0.011**	0.011	0.008
	(1.359)	(1.356)	(2.755)	(2.089)	(1.267)	(1.388)
GP	0.009	0.010	0.012^{*}	0.008	0.006	0.008
	(1.552)	(1.612)	(1.838)	(1.237)	(1.086)	(1.319)
OP	0.030***	0.030***	0.033***	0.031***	0.022***	0.027***
	(3.925)	(3.975)	(3.896)	(3.452)	(3.522)	(3.700)
OA	0.001	0.001	0.001	0.001	0.000	0.000
	(0.322)	(0.431)	(0.388)	(0.224)	(0.002)	(-0.089)
ТА	-0.008^{**}	-0.008^{**}	-0.004	-0.005	-0.005	-0.006
	(-2.128)	(-2.184)	(-0.906)	(-1.388)	(-1.407)	(-1.388)
AG	-0.011^{**}	-0.011^{**}	-0.012***	-0.010^{**}	-0.008^{**}	-0.007^{*}
	(-2.373)	(-2.439)	(-2.633)	(-2.248)	(-2.026)	(-1.667)
Adj. R^2	0.045	0.038	0.032	0.028	0.030	0.036
obs.	240,844	240,844	240,844	240,844	240,844	240,844
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Table 6. CNN Buy Probability, Firm Characteristics, and Historical Earnings. The table presents results of quarterly weighted Fama and MacBeth (1973) regressions in the out-of-sample period (1994Q3-2023Q2, 116 quarters) using the CNN buy probability (CNNBP) as the dependent variable. The weights correspond to the number of observations used in The independent variables in specification 1 are each quarterly cross-sectional regression. standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), market capitalization (SIZE), book-to-market ratio (BM), earnings announcement return ($\operatorname{RET}[-1, -1]$), pre-announcement return ($\operatorname{RET}[-30, -2]$), earnings persistence (PERSIST), earnings volatility (VOL), gross profitability (GP), operating profitability (OP), total accruals (TA), operating accruals (OA), and asset growth (AG). The independent variables in specification 2 are the return on assets (ROA) in the most recent eight quarters (quarter q - 7 to q). All variables are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of zero. See Appendix for variable definitions. Newey and West (1987) t-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

	(1) CNNBP		(2)
	CNNBP		CNNBP
Intercept	-0.002	Intercept	-0.004^{***}
	(-1.598)		(-3.040)
SUE	0.510***	ROA_q	0.644***
	(77.045)	-	(41.254)
\mathbf{EA}	0.040***	ROA_{q-1}	0.038***
	(6.851)	-	(6.965)
TREND	0.004	ROA_{q-2}	-0.073^{***}
	(0.792)	-	(-14.535)
$\operatorname{RET}[-1,1]$	0.050***	ROA_{q-3}	-0.333^{***}
	(20.524)	-	(-62.953)
RET[-30, -2]	0.024***	ROA_{q-4}	-0.428^{***}
	(7.350)	-	(-48.276)
PERSIST	-0.044^{***}	ROA_{q-5}	0.079***
	(-8.605)		(16.625)
VOL	0.246^{***}	ROA_{q-6}	0.044***
	(21.668)		(12.322)
SIZE	-0.274^{***}	ROA_{q-7}	-0.021^{***}
	(-18.569)		(-4.567)
BM	0.037***		
	(5.876)		
GP	0.010**		
	(2.155)		
OP	-0.023^{***}		
	(-3.938)		
OA	0.001		
	(0.186)		
ТА	-0.019^{***}		
	(-8.196)		
AG	0.047^{***}		
	(9.380)		
Adj. R^2	0.328	Adj. R^2	0.300
obs.	240,844	obs.	240,043
		0.55.	= 10,0 10

Table 7. CNN Buy Probability, Future Earnings Growth, and Future Three-Day Abnormal Returns around Earnings Announcements. The table presents results of quarterly weighted Fama and MacBeth (1973) regressions in the out-of-sample period (1994Q3-2023Q2, 116 quarters) using one-quarter-ahead standardized unexpected earnings (SUE_{q+1}) or the three-day abnormal return around the next earnings announcement date $(\text{RET}[-1,1]_{q+1})$ as the dependent variable. The weights correspond to the number of observations used in each quarterly cross-sectional regression. The independent variables include CNN buy probability (CNNBP), standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), market capitalization (SIZE), book-to-market ratio (BM), earnings announcement return ($\operatorname{RET}[-1, -1]$), pre-announcement return ($\operatorname{RET}[-30, -2]$), earnings persistence (PERSIST), earnings volatility (VOL), gross profitability (GP), operating profitability (OP), total accruals (TA), operating accruals (OA), and asset growth (AG). All variables are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of zero. See Appendix for variable definitions. Newey and West (1987) t-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

	$(1) \\ \mathrm{SUE}_{q+1}$	$(2) \\ \operatorname{SUE}_{q+1}$	(3) RET $[-1,1]_{q+1}$	(4) RET $[-1,1]_{q+1}$
			$\frac{1001[-1,1]q+1}{0.002^{***}}$	
Intercept	$0.002 \\ (0.779)$	$0.002 \\ (0.802)$	(3.564)	0.002^{***} (3.586)
CNNBP	$(0.113)^{(0.113)}$	(0.802) 0.133^{***}	(3.304) 0.004^{***}	(3.380) 0.004^{***}
UNINDI	(13.731)	(20.601)	(4.282)	(5.185)
SUE	0.377***	0.346^{***}	-0.002	-0.002
DOL	(30.493)	(31.877)	(-1.319)	(-1.531)
$\mathbf{E}\mathbf{A}$	-0.090^{***}	-0.086^{***}	0.001	0.001
200	(-24.240)	(-24.216)	(0.866)	(1.088)
TREND	0.021***	0.024***	0.000	0.001
-	(4.331)	(6.614)	(0.165)	(1.470)
RET[-1,1]	0.078^{***}	0.077^{***}	0.004***	0.004***
	(37.487)	(37.871)	(4.348)	(3.899)
RET[-30, -2]	0.067^{***}	0.066***	0.001	0.000
	(16.162)	(16.506)	(0.548)	(-0.110)
PERSIST	0.001	-0.002	-0.002^{***}	-0.001^{**}
	(0.138)	(-0.510)	(-3.011)	(-2.029)
VOL	0.009**	-0.076^{***}	-0.001	-0.002
	(2.237)	(-8.943)	(-0.904)	(-1.283)
SIZE		0.119***		0.003
		(10.848)		(1.253)
BM		-0.018***		0.007***
a 5		(-3.278)		(5.733)
GP		0.012**		0.006***
ÔÐ		(2.107)		(4.640)
OP		-0.022^{***}		0.005^{***}
OA		$(-3.443) \\ -0.016^{***}$		(4.847)
0A				0.000
ТА		$(-5.093) \\ -0.002$		$(0.521) \\ -0.001$
IA		(-0.728)		(-1.326)
AG		(-0.728) -0.015^{***}		(-1.520) -0.002^{**}
AO		(-3.934)		(-2.187)
Adj. R^2	0.214	0.228	0.003	0.007
obs.	237,118	237,118	236,239	236,239
	,			

Table 8. Test of Market Efficiency for the *CNN Buy* Features Effect. This table reports the regression results from nonlinear generalized least squares estimation of the following two equations in the out-of-sample period (1994Q3-2023Q2, 116 quarters)

Forecasting equation:
$$\text{SUE}_{q+1} = \alpha + \gamma_1 \text{CNNBP}_q + \sum \gamma_c \text{Controls}_q + \delta_{q+1}$$

Pricing equation: $\text{AR}_{q+1} = \beta(\text{SUE}_{q+1} - \alpha^* - \gamma_1^* \text{CNNBP}_q - \sum \gamma_c^* \text{Controls}_q) + \varepsilon_{q+1}$.

The control variables include standardized unexpected earnings (SUE), earnings acceleration (EA), trend in gross profitability (TREND), market capitalization (SIZE), book-to-market ratio (BM), earnings announcement return (RET[-1,-1]), pre-announcement return (RET[-30,-2]), earnings persistence (PERSIST), earnings volatility (VOL), gross profitability (GP), operating profitability (OP), total accruals (TA), operating accruals (OA), and asset growth (AG). AR is the abnormal return from a 3-day window around quarter q + 1's earnings or the quarter-long window starting two days after the quarter q earnings and ending on the next announcement date. All variables except for the abnormal return AR are converted into scaled ranks ranging from -0.5 to 0.5 with a mean of zero. t-statistics based on firm and quarter double-clustered standard errors are reported in parentheses. The likelihood ratio statistic for testing $\gamma_1 = \gamma_1^*$ is distributed asymptotically as $\chi^2(1)$. ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

Parameter	3-day	Quarter-long
γ_1	0.136^{***} (19.783)	0.136^{***} (19.816)
γ_1^*	0.062^{***} (3.814)	0.010 (0.233)
β	0.062^{***} (38.571)	0.010^{***} (18.547)
Market efficiency test $(\gamma_1 = \gamma_1^*)$	3-day	Quarter-long
Likelihood ratio statistic	8.218***	21.444***

Table 9. Month-Based Rebalancing Strategy. At the end of each month quarters), we sort firms into deciles based on their most recent CNN buy probabilit earnings). For a firm to enter the portfolio formation at the end of month t , we earnings to be within three months prior to portfolio formation. This table reports returns, as well as factor-adjusted returns in the subsequent month $t + 1$ for each t -statistics with three lags are reported in parentheses, and ***, **, and * indicate s	conth-Base sort firms ir a firm to within three all as factor- th three lags	Month-Based Rebalancing Strategy. At the we sort firms into deciles based on their most recent For a firm to enter the portfolio formation at the be within three months prior to portfolio formation. well as factor-adjusted returns in the subsequent m with three lags are reported in parentheses, and ***,	cing Strategy. used on their mo- ortfolio formatio for to portfolio f urns in the subse l in parentheses,	Strategy. At the a their most recent o formation at the portfolio formation. In the subsequent more the the subsequent more the set.	end of each CNN buy pr end of mont end of mont this table 1 onth $t + 1$ fc **, and * in	month t in robability (c th t , we rec reports the i or each CNN dicate signif	the out-o omputed u puire that a average equ V buy prob icance at th	At the end of each month t in the out-of-sample period (1994Q3-2023Q2, 116 t recent CNN buy probability (computed using firms' most recent eight quarterly at the end of month t , we require that announcement date of its most recent mation. This table reports the average equal-weighted and value-weighted excess quent month $t + 1$ for each CNN buy probability decile. Newey and West (1987) and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively	od (1994Q3 ost recent ei t date of its und value-we Newey and d 10% level	end of each month t in the out-of-sample period (1994Q3-2023Q2, 116 <i>CNN buy</i> probability (computed using firms' most recent eight quarterly end of month t , we require that announcement date of its most recent This table reports the average equal-weighted and value-weighted excess onth $t + 1$ for each <i>CNN buy</i> probability decile. Newey and West (1987) **, and * indicate significance at the 1%, 5%, and 10% level, respectively.
I		Equal-	Equal-weighted returns	urns			Value	Value-weighted returns	urns	
CNNBP	Excess	FF4	FF6	HMXZ5	DHS3	Excess	FF4	FF6	HMXZ5	DHS3
deciles	return	alpha	alpha	alpha	alpha	return	alpha	alpha	alpha	alpha
Low	0.006^{*}	-0.001	-0.002^{***}	-0.002	0.000	0.006^{**}	0.000	-0.001	-0.001	-0.001
	(1.955)	(-0.868)	(-2.681) ((-1.289)	(0.041)	(2.412) ((-0.189)	(-1.445)	(-0.765)	(-0.725)
2	0.007^{**}	0.000	-0.001	0.000	0.001	0.007^{***}	0.001	0.000	0.000	0.000
	(2.289)	(0.045)	(-0.827)	(0.166)	(0.914)	(3.003)	(1.018)	(-0.268)	(-0.025)	(0.253)
3	0.009^{**}	0.002^{*}	0.002^{*}	0.003^{**}	0.004^{**}	0.007^{***}	0.001	0.000	0.001	0.000
	(2.546)	(1.881)	(1.667)	(2.056)	(2.121)	(3.093)	(0.591)	(-0.285)	(0.501)	(0.058)
4	0.009^{**}	0.002	0.002^{*}	0.004^{**}	0.005^{**}	0.007^{***}	0.000	0.000	0.000	-0.001
	(2.336)	(1.331)	(1.728)	(2.105)	(2.019)	(2.693)	(0.118)	(-0.195)	(-0.286)	(-0.462)
ъ	0.011^{***}	0.003^{**}	0.003^{***}	0.005^{***}	0.006^{**}	0.008^{***}	0.001	0.001	0.000	0.001
	(2.811)	(2.560)	(2.674)	(3.248)	(2.586)	(3.344)	(1.263)	(0.681)	(0.111)	(1.052)
6	0.011^{***}	0.004^{***}	0.004^{***}	0.005^{***}	0.006^{**}	0.009^{***}	0.002	0.002	0.002	0.003^{**}
	(2.872)	(2.958)	(3.148)	(2.789)	(2.531)	(3.183)	(1.549)	(1.251)	(1.337)	(2.058)
7	0.012^{***}	0.004^{***}	0.003^{***}	0.005^{***}	0.006^{***}	0.007^{***}	0.000	-0.001	-0.002	0.001
	(3.223)	(3.615)	(3.182)	(3.285)	(3.025)	(2.826)	(0.162)	(-0.491)	(-1.191)	(0.771)
x	0.013^{***}	0.005^{***}	0.004^{***}	0.005^{***}	0.006^{***}	0.008^{***}	0.000	0.000	0.001	0.000
	(3.335)	(3.897)	(3.788)	(3.867)	(2.948)	(2.699)	(0.175)	(0.364)	(0.721)	(0.096)
6	0.015^{***}	0.006^{***}	0.005^{***}	0.006^{***}	0.008^{***}	0.011^{***}	0.003^{**}	0.004^{***}	0.003^{**}	0.005^{***}
	(4.081)	(5.032)	(4.604)	(5.165)	(4.082)	(3.793)	(2.198)	(2.798)	(2.268)	(3.077)
High	0.016^{***}	0.006^{***}	0.006^{***}	0.007^{***}	0.009^{***}	0.012^{***}	0.003^{**}	0.003^{**}	0.004^{**}	0.004^{***}
	(4.518)	(4.689)	(4.620)	(4.458)	(4.461)	(3.758)	(2.070)	(2.467)	(2.373)	(3.417)
High-Low	0.010^{***}	0.007^{***}	0.008^{***}	0.008^{***}	0.009^{***}	0.006^{***}	0.003^{*}	0.005^{***}	0.004^{**}	0.005^{***}
	(6.055)	(5.284)	(6.237)	(5.274)	(6.298)	(3.115)	(1.852)	(2.911)	(2.294)	(3.155)

Table 10. Alternative Modeling Choices in Training. This table reports the last row of Table 2 under different modeling choices in training. The baseline modeling choices in Table 2 are: two-dimensional CNN model; 3 blocks; filter size = $(7 \times 7, 3 \times 3, 3 \times 3)$; dropout rate = 0.5; batch normalization = Yes; Xavier initialization = Yes; activation = Leaky ReLU. See Table 2 and Section 6.2 for more details. Newey and West (1987) *t*-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

	Difference in	the 63-day Bl	HAR between	the highest an	Difference in the 63-day BHAR between the highest and lowest CNNBP deciles	3P deciles
Alternative modeling choices	MAR	SAR	FF4	FF6	HMXZ5	DHS3
3 blocks; filter size = $(5 \times 5, 3 \times 3, 3 \times 3)$	0.034^{***}	0.034^{***}	0.029^{***}	0.031^{***}	0.034^{***}	0.033^{***}
	(6.537)	(6.456)	(7.255)	(7.677)	(7.582)	(7.528)
3 blocks; filter size = $(3 \times 3, 3 \times 3, 3 \times 3)$	0.035^{***}	0.035^{***}	0.031^{***}	0.031^{***}	0.035^{***}	0.034^{***}
	(7.414)	(7.367)	(8.516)	(9.154)	(8.932)	(8.678)
2 blocks; filter size = $(i \times i, 3 \times 3)$	(7539)	0.032*** (7 495)	0.027*** (7 300)	0.029*** (7 971)	(8.371)	0.031*** (7 850)
2 blocks; filter size = $(5 \times 5, 3 \times 3)$	0.035^{***}	0.034^{***}	0.029^{***}	0.030^{***}	0.032^{***}	0.033^{***}
	(7.910)	(7.802)	(7.752)	(8.560)	(8.879)	(9.057)
2 blocks; filter size = $(3 \times 3, 3 \times 3)$	0.034^{***}	0.033^{***}	0.029^{***}	0.030^{***}	0.032^{***}	0.032^{***}
	(7.715)	(7.506)	(8.312)	(8.871)	(8.687)	(9.005)
Dropout rate $= 0$	0.032^{***}	0.031^{***}	0.027^{***}	0.027^{***}	0.030^{***}	0.031^{***}
	(6.215)	(6.182)	(6.600)	(7.050)	(6.446)	(6.967)
Batch normalization $=$ no	0.034^{***}	0.033^{***}	0.029^{***}	0.030^{***}	0.032^{***}	0.032^{***}
	(7.425)	(7.205)	(8.400)	(9.328)	(8.491)	(8.723)
Xavier initialization $=$ no	0.035^{***}	0.034^{***}	0.030^{***}	0.031^{***}	0.034^{***}	0.033^{***}
	(6.433)	(6.372)	(7.476)	(7.917)	(7.539)	(7.394)
Activation = ReLU	0.035^{***}	0.034^{***}	0.031^{***}	0.032^{***}	0.033^{***}	0.033^{***}
	(7.343)	(7.162)	(8.594)	(9.466)	(8.164)	(8.050)
Panel B: One-dimensional CNN model						
	Difference in	the 63-day Bl	HAR between	the highest an	Difference in the 63-day BHAR between the highest and lowest CNNBP deciles	3P deciles
Alternative modeling choices	MAR	SAR	FF4	FF6	HMXZ5	DHS3
3 blocks; filter size = $(1 \times 7, 1 \times 3, 1 \times 3)$	0.007*	0.007^{*}	0.009^{**}	0.009^{**}	0.010^{***}	0.010^{***}
	(1.905)	(1.926)	(2.305)	(2.454)	(2.663)	(2.751)
3 blocks; filter size = $(1 \times 5, 1 \times 3, 1 \times 3)$	0.006	0.005	0.005	0.006*	0.006^{*}	0.007^{**}
	(1.504)	(1.475)	(1.420)	(1.660)	(1.676)	(2.012)
3 blocks; filter size = $(1 \times 3, 1 \times 3, 1 \times 3)$	0.002	0.001	0.003	0.003	0.001	0.003
	(0.773)	(0.567)	(1.268)	(1.140)	(0.359)	(0.972)
2 blocks; filter size = $(1 \times 7, 1 \times 3)$	0.001	(0.000)	0.001	(0.001)	0.002	0.002
9 blocks: filter size $-(1 \times 5, 1 \times 3)$	(0.430) 0.001	(0.013) 0.001	(0.210)	(0.230) 0.009	(0.349)	(U.839) 0.000
<	(0.311)	(0.234)	(-0.508)	(-0.574)	(0.329)	(0.041)
2 blocks; filter size = $(1 \times 3, 1 \times 3)$	0.004	0.005	0.005	0.004	0.008^{*}	0.007
	(0.901)	(1.049)	(1.084)	(0.937)	(1.738)	(1.571)

Table IA1. CNN Buy Probability and Post-Earnings Announcement Drift: Double Sorts. This table reports the average 63-day buy-and-hold market-adjusted return (MAR) after earnings announcements for portfolios formed based on the CNN buy probability quintiles and one of the six firm characteristics including pre-announcement return (RET[-30, -2]), earnings persistence (PERSIST), earnings volatility (VOL), market capitalization (SIZE), operating profitability (OP), total accruals (TA), and operating accruals (OA) using independent two-way sorting in the out-of-sample period (1994Q3-2023Q2, 116 quarters). The quintile cutoffs are based on the distribution of these variables in the previous quarter. See Appendix for variable definitions. Newey and West (1987) *t*-statistics with three lags are reported in parentheses, and ***, **, and * indicate significance at the 1%, 5%, and 10% level, respectively.

Panel A: Two-way sortin	ng, controlling	g for pre-ann	ouncement re	turn ($\operatorname{RET}[-3]$	(30, -2])	
		(CNN buy prob	ability quintil	es	
$\operatorname{RET}[-30, -2]$ quintiles	Low	2	3	4	High	High-Low
Low	0.006	0.009	0.013	0.016	0.036***	0.030***
2	$(0.628) \\ 0.000$	$(0.849) \\ 0.004$	$(1.057) \\ 0.009$	$(1.393) \\ 0.016^{**}$	$(3.880) \\ 0.020^{***}$	$(5.139) \\ 0.020^{***}$
3	$(-0.030) \\ -0.004$	$(0.676) \\ 0.004$	$(1.332) \\ 0.004$	$(2.385) \\ 0.012^*$	$(3.404) \\ 0.022^{***}$	$(5.544) \\ 0.026^{***}$
პ	(-0.811)	(0.813)	(0.691)	(1.677)	(4.187)	(6.515)
4	(-0.009^{*}) (-1.750)	(0.003) (0.413)	0.008 (1.115)	0.014^{**} (2.189)	(0.023^{***})	(0.031^{***})
High	-0.016^{**}	-0.005	(1.113) 0.011	0.014	0.036^{***}	0.052***
	(-2.203)	(-0.551)	(1.145)	(1.536)	(4.102)	(8.324)
High-Low	-0.022^{***} (-2.659)	-0.015^{*} (-1.781)	$-0.002 \\ (-0.216)$	$-0.002 \\ (-0.313)$	$\begin{array}{c} 0.000 \\ (-0.047) \end{array}$	
	. 11.	с ·	•			

Panel B: Two-way sorting, controlling for earnings persistence (PERSIST)

		C	<i>CNN buy</i> prob	ability quintil	es	
PERSIST quintiles	Low	2	3	4	High	High-Low
Low	-0.005	0.001	0.005	0.012	0.029***	0.034***
2	$(-1.033) \\ 0.000$	$(0.204) \\ 0.008$	$(0.636) \\ 0.015^*$	$(1.574) \\ 0.015^*$	$(4.461) \\ 0.026^{***}$	$(5.995) \\ 0.026^{***}$
3	$(-0.025) \\ -0.005$	$(1.067) \\ 0.009$	$(1.946) \\ 0.009$	$(1.902) \\ 0.017^{**}$	$(4.749) \\ 0.029^{***}$	$(5.770) \\ 0.034^{***}$
4	$(-0.970) \\ -0.006$	$(1.205) \\ 0.003$	$(1.128) \\ 0.008$	$(2.091) \\ 0.015^{**}$	$(4.515) \\ 0.025^{***}$	$(7.839) \\ 0.031^{***}$
High	$(-1.052) \\ -0.004$	$(0.277) \\ 0.000$	$(1.038) \\ 0.006$	$(2.081) \\ 0.009$	$(3.918) \\ 0.023^{***}$	$(5.331) \\ 0.027^{***}$
	(-0.738)	(-0.065)	(0.661)	(1.190)	(3.239)	(4.647)
High-Low	$\begin{array}{c} 0.001 \\ (0.360) \end{array}$	$-0.002 \\ (-0.424)$	$\begin{array}{c} 0.001 \\ (0.309) \end{array}$	$-0.003 \\ (-0.586)$	$-0.006 \\ (-0.983)$	

Panel C: Two-way sorting, controlling for earnings volatility (VOL)

		C	NN buy prob	ability quintil	es	
VOL quintiles	Low	2	3	4	High	High-Low
Low	-0.005	0.001	0.005	0.012	0.029***	0.034***
2	(-1.033) 0.000 (-0.025)	(0.204) 0.008 (1.067)	$(0.636) \\ 0.015^{*} \\ (1.046)$	(1.574) 0.015^{*}	(4.461) 0.026^{***}	(5.995) 0.026^{***}
3	(-0.025) -0.005	(1.067) 0.009 (1.225)	$(1.946) \\ 0.009 \\ (1.122)$	(1.902) 0.017^{**}	(4.749) 0.029^{***}	(5.770) 0.034^{***}
4	(-0.970) -0.006	(1.205) 0.003 (0.277)	$(1.128) \\ 0.008 \\ (1.020)$	(2.091) 0.015^{**}	(4.515) 0.025^{***}	(7.839) 0.031^{***}
High	$\begin{array}{c} (-1.052) \\ -0.004 \\ (-0.738) \end{array}$	$(0.277) \\ 0.000 \\ (-0.065)$	$\begin{array}{c}(1.038)\\0.006\\(0.661)\end{array}$	$(2.081) \\ 0.009 \\ (1.190)$	$(3.918) \\ 0.023^{***} \\ (3.239)$	$(5.331) \\ 0.027^{***} \\ (4.647)$
High-Low	$\begin{array}{c} 0.001 \\ (0.360) \end{array}$	$-0.002 \\ (-0.424)$	$\begin{array}{c} 0.001 \\ (0.309) \end{array}$	$-0.003 \\ (-0.586)$	$-0.006 \ (-0.983)$	

Panel D: Two-w	vay sorting, co	ntrolling for n	narket capitali	zation (SIZE)		
			CNN buy prob	pability quintile	es	
SIZE quintiles	Low	2	3	4	High	High-Low
Low	-0.018*	0.009	0.012	0.020*	0.056***	0.074***
	(-1.945)	(0.613)	(0.998)	(1.837)	(5.679)	(13.846)
2	-0.006	0.005	0.012	0.018^{*}	0.043***	0.049**
	(-0.939)	(0.518)	(1.179)	(1.698)	(4.934)	(8.833)
3	-0.002	0.006	0.013	0.017^{*}	0.022***	0.024**
	(-0.320)	(0.828)	(1.598)	(1.910)	(3.338)	(5.235)
4	0.001	0.003	0.003	0.006	0.011**	0.010**
	(0.100)	(0.575)	(0.685)	(1.201)	(2.338)	(2.508)
High	0.000	-0.001	[0.002]	0.004	0.009**	0.009^{*}
	(-0.103)	(-0.258)	(0.696)	(1.365)	(2.203)	(1.840)
High-Low	0.017**	-0.010	-0.010	-0.016^{*}	-0.047^{***}	
0	(2.110)	(-0.691)	(-0.887)	(-1.682)	(-4.815)	
Panel E: Two-w	ay sorting, co	ntrolling for o	perating profit	ability (OP)		
			CNN buy prob	pability quintile	es	
OP quintiles	Low	2	3	4	High	High-Low
Low	-0.017	-0.001	0.002	0.001	0.032***	0.049**
	(-1.474)	(-0.111)	(0.167)	(0.102)	(3.234)	(6.128)
2	-0.008^{\prime}	0.007	0.009^{\prime}	0.015^{*}	0.030^{***}	0.037**
	(-1.268)	(0.886)	(1.199)	(1.677)	(3.628)	(6.455)
3	-0.005	$-0.003^{'}$	0.016^{*}	0.019^{**}	0.028***	0.033**
	(-0.887)	(-0.492)	(1.952)	(2.503)	(4.795)	(8.142)
4	-0.004	0.007	0.008	`0.016 ^{**}	0.024***	0.028**
	(0.707)	(1, 000)	(1, 0, 1, 0)	(0 550)	(0 = 0 1)	(F 14F)

Table IA1.CNN Buy Probability and Post-Earnings Announcement Drift: Double Sorts(continued).

Panel F: Two-way	sorting.	controlling :	for	operating	accruals ((OA)	

(1.220)

0.010*

(1.762)

0.012

(1.082)

(-0.787)

0.004

(0.810)

(2.017)

0.021**

High

High-Low

	CNN buy probability quintiles					
OA quintiles	Low	2	3	4	High	High-Low
Low	-0.008	0.004	0.009	0.014	0.033***	0.041***
2	$(-1.071) \\ 0.001$	$(0.363) \\ 0.006$	$(0.910) \\ 0.010$	$(1.376) \\ 0.014^*$	$(4.291) \\ 0.027^{***}$	$(7.451) \\ 0.026^{***}$
3	$(0.237) \\ -0.004$	$(0.846) \\ 0.004$	$(1.325) \\ 0.015^*$	$(1.909) \\ 0.018^{**}$	$(4.200) \\ 0.025^{***}$	$(4.839) \\ 0.029^{***}$
4	$(-0.735) \\ -0.004$	$(0.752) \\ 0.007$	$(1.896) \\ 0.006$	$(2.195) \\ 0.015^{**}$	$(3.969) \\ 0.027^{***}$	$(7.300) \\ 0.031^{***}$
High	$(-0.820) \\ -0.007 \\ (-1.184)$	$(0.975) \\ -0.003 \\ (-0.362)$	$(1.055) \\ 0.004 \\ (0.452)$	$(2.560) \\ 0.009 \\ (1.235)$	$(4.901) \\ 0.027^{***} \\ (4.195)$	$(7.319) \\ 0.034^{***} \\ (7.041)$
High-Low	$\begin{array}{r} (1.104) \\ \hline 0.001 \\ (0.176) \end{array}$	$\frac{-0.007}{(-0.848)}$	$\frac{-0.006}{(-1.038)}$	$-0.006 \\ (-0.813)$		(1.041)

(1.218)

 0.012^{*}

(1.887)

0.010

(1.288)

(2.552) 0.021^{***}

0.020**

(3.406)

(2.468)

 $(3.784) \\ 0.023^{***}$

(3.486)

-0.008

(-0.963)

(5.145) 0.019^{***}

(3.438)

Panel G: Two-way sorting, controlling for total accruals (TA)							
	CNN buy probability quintiles						
TA quintiles	Low	2	3	4	High	High-Low	
Low	-0.008 (-0.894)	$0.012 \\ (0.910)$	0.009 (0.828)	$0.009 \\ (0.831)$	0.037^{***} (4.012)	0.045^{***} (6.271)	
2	(-0.003) (-0.415)	(-0.001) (-0.073)	(0.011) (1.299)	0.017^{**} (2.207)	$(1.021)^{***}$ (4.606)	0.033^{***} (7.087)	
3	(-0.002) (-0.335)	0.002 (0.333)	(0.009) (1.448)	(0.015^{**}) (2.073)	(0.027^{***}) (3.879)	0.029^{***} (6.916)	
4	-0.003 (-0.656)	(0.004)	0.011^{*} (1.781)	0.021^{***} (3.112)	(0.027^{***})	0.030^{***} (6.644)	
High	(-0.005) (-0.956)	(-0.002) (-0.348)	$egin{array}{c} 0.004 \ (0.538) \end{array}$	(0.008) (1.282)	(0.017^{***}) (2.833)	(0.022^{***}) (4.321)	
High-Low	$\begin{array}{c} 0.003 \ (0.375) \end{array}$	-0.014 (-1.449)	$-0.005 \ (-0.811)$	$-0.001 \\ (-0.169)$	-0.021^{***} (-3.186)		

Table IA1.CNN Buy Probability and Post-Earnings Announcement Drift: Double Sorts(continued).

References

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